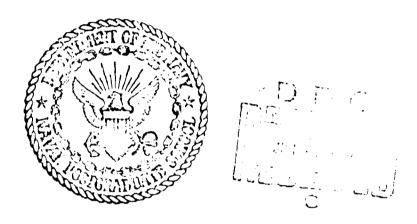
NAVAL POSTGRADUATE SCHOOL Monterey, California



THESIS

OPTIMUM THRESHOLD DETECTION FOR AN INFRARED NUTATING DETECTION SYSTEM

by

William Earl Major, II

Thesis Advisor:

Harold Titus

March 1973

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A threshold detection system based on the Neyman-Pearson criterion is derived for an infrared nutating optical system. Detection is optimum for small signalto-noise ratios and a Gaussian uncertainty in the pointing error of the optical system. The signal spectrum and background power spectral density for a rectangular space filter are computed numerically and used to specify the matched filter for the threshold detection system and evaluate its performance.

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Optimum Threshold Detection

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Infrared Natating Detection System

bу

William Earl Major, II Lieutenant, United States Navy B.S., Auburn University, 1965

Submitted in partial fulfillment of the requirements for the degree of

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1. INTRODUCTION

An infrared (IR) detection system determines the presence or absence of a target in a particular area of space by processing the received IR energy from that area. The target must be found amid ever-present background radiance and any internal noise generated by the detection system. In this thesis, one type of system, called a nutating optical system, is considered.

Infrared detection is a combination of spatial frequency processing and temporal processing. Spatial frequency processing is a means of achieving a modulated output from the IR radiance passing into the optical system so that information may be easily extracted [Ref. 14]. To extract this information, some form of temporal processing is used. In a nutating optical system, the optical axis rotates about an axis perpendicular to the image plane in such a way that a point image traces out a circle in the image plane. spatial frequency filter is a piece of IR sensitive material in the image plane which produces a voltage output that is amplified and fed into a temporal filter for processing. A nutating optical system is frequently used in a missile seeker because it produces a signal when there is no tracking error and the rotating optics can be made into a gyro that provides the inertial reference.

This work is a synthesis and extension of two recently published articles in the literature. The first, by

Harger [Ref. 3], formulates the detection theory of a known signal in background and white noise. Since in 1R target detection the signal is never known completely, llarger's work is extended to the case where the target amplitude and position are unknown. However, a known target shape is assumed as well as a priori statistics of amplitude and position. Since detection is most difficult when the amplitude is small, a test (called the threshold detector) is derived that is optimum for weak signals.

To specify the form of the threshold detector and to calculate it: performance, an eigenvalue integral equation must be solved. Ordinarily, this is a difficult undertaking but the solution becomes trivial if the covariance function of the background process is periodic. This is exactly the condition realized for the background process out of a nutating detector provided that the phase, which is physically unimportant, is averaged out.

Samuelsson's results [Ref. 11] specify the form of the signal and background coefficients in terms of the system parameters and signal and background radiance. These derivations were first published in a Swedish internal report [Ref. 10], then re-derived in English by Yoshitani [Ref. 6]. The latter's derivations are included in the appendices.

Samuelsson's equations have been applied to a rectangular IR detector assuming a Wiener spectrum for the random background and the coefficients have been calculated.

Using these coefficients, the optimum filter has been determined and the probability of detection computed as a function of signal-to-noise ratios and background-to-noise ratios.

II. SPATIAL PROCESSING

The spatial frequency filtering portion of the detection system is comprised of a lens system and an IR sensitive detected located in the image plane, the focal plane of the optics. To adequately derive information from the IR radiation, the lens system and detector must be of a nature that causes the veltage output to be modulated in a way that information about a target within the field of view can be processed. In a practical nutating optical system, the lens system moves and the detector is stationary. However, for conceptual purposes, this is equivalent to a stationary lens system and nutating detector.

A. TARGET SPECTRUM

The object to be detected, the target, is located in the object plane a distance R from the lens and parallel to the image plane as shown in Figure 1.1. The angle from the perpendicular axis connecting the planes to the target coordinates measured in the x-direction is called x and is measured in milliradians; the angle in the y-direction is y. The target is considered to be close to the optical axis and at a sufficient distance, R, such that small angle approximations may be made.

$$x = \tan x$$
 $y = \tan y$
(2.1)

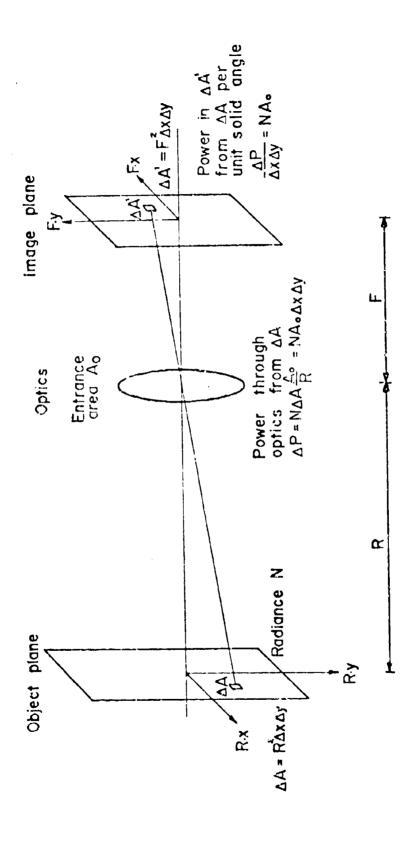


Figure 1.1. Optical Imaging.

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Consider a target with a radiance distribution N(r) (W/cm²-sr) which is the radiant power into a unit solid angle per unit projected area of the sources. r = r(x,y) is a vector whose origin is the perpendicular axis connecting the coordinates.

Assuming no atmospheric attenuation (this can be included in $N(\underline{r})$ if it exists) and perfect optics, the power distribution in the image plane is

$$P_{i}(r) = N(r) \cdot A_{0} (w/sr)$$
 (2.2)

where A_0 is the effective entrance area of the optics.

Dividing by A_0 , one can see that the radiance distribution in the image plane is numerically equal to the radiance distribution in the object plane.

However, because the optics is not perfect, a point object given by

$$Hp \delta(r) = Hp \delta(x) \delta(y) (w) , \qquad (2.3)$$

(where Hp (W/cm^2) is the irradiance at the optics received from the point source) is imaged as a power distribution

$$Hp \cdot f_o(r) \qquad (W/cm^2 Sr) \qquad (2.4)$$

where $f_0(r)$ (sr⁻¹) is the point spread function of the optics. Irradiance is defined as the radiant flux indicent on a surface of unit area [Ref. 5].

Therefore, the radiance distribution, N(r), in the object plane will be imaged as

$$N'(r) = \int N(r-s) f_0(s) d^2s (W/cm^2Sr)$$
 (2.5)

where $d^2s = dxdy$ and the integration is over the entire image plane.

The detector is a space filter in the image plane which may be a moving retical or simply an aperture across which the power is integrated.

The power in the image plane incident on the detector is

$$H = \int N'(r) \tau(r) d^2r \qquad (W/cm^2) \qquad (2.6)$$

where $\tau(r) = \tau(x,y)$ is the transmittance of the detector.

When the detector moves as a function of time, the power incident on the detector becomes a function of time,

$$H(t) = \int N'(\tau) \tau(\tau - \varrho(t)) d^2\tau \qquad (W/cm^2) \qquad (2.7)$$

where $\rho(t)$ describes the movement of the detector in the image plane.

A nutating detector moves circularly in the image plane but its orientation remains fixed, i.e., the coordin te system (x',y') shown in Fig. 1.2 does not rotate. The nutation radius measured with respect to the optical axis is given by ρ .

Assuming a nutating optics and no relative motion between the target and the optics, the radiance on the detector will be periodic and thus can be expanded in a Fourier series

$$H(t) = \sum_{n=0}^{\infty} H_n e^{jn\omega_0 t} \qquad 0 \le t \le T$$
 (2.8)

where

$$H_n = \frac{1}{T} \int_0^T H(t) e^{-jn\omega_0 t} dt$$
.

Replacing H(t) by equation (2.7),

$$H_{n} = \frac{1}{T} \int_{0}^{T} dt e^{-jn\omega_{0}t} d^{2}r N'(r) (r-\rho(t)) \qquad (2.9)$$

The motion describing function, $\varrho(t)$, for circular nutation is

$$\rho(t) = \rho(\cos \omega_0 t, \sin \omega_0 t) , \qquad (2.10)$$

where ω_0 = 2 π/T (rad/sec) is the radian frequency of nutation and T is the period for one nutation.

Substituting (2.10) in (2.9), ${\rm H}_{\pi}$ can be shown (see Appendix A) to be

$$H_{n} = j^{n} \int N'(k) \tau^{*}(k) e^{-jn\phi} J_{n}(2\pi\rho \sqrt{k_{x}^{2} + k_{y}^{2}}) d^{2}k \qquad (2.11)$$

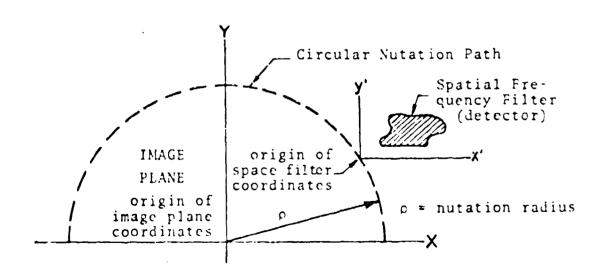


Figure 1.2. Circular Nutation in the Image Plane.

where $k = (k_x, k_y)$, two-dimensional spacial Fourier transform coordinates [Ref. 8],

- N'(k) = transform of radiance distribution in the
 image plane,

$$\phi = \tan^{-1} k_y/k_x,$$

 ρ = nutation radius,

and $J_n(z)$ is the nth order Bessel function of the first kind. It has been assumed that the image plane is very large compared to the radiance distribution from a target so that the image plane may be considered infinite for mathematical purposes.

B. CORRELATION OF BACKGROUND NOISE

The background power incident on the detector is from a sample background scene where a scene is a two-dimensional random process characterized by the radius vector $\mathbf{r} = (x,y)$ from the optical axis.

Let N(r) be a sample scene radiance distribution on the object plane from an ensemble of scenes which has been given a suitable probability structure. The background power incident on the detector is

$$B(t) = \int N'(r) \tau(r-\rho(t)) d^2r$$
 (2.12)

The covariance function is defined by taking the expected value

$$E\{[B(t)\} - E\{B(t)\}\}[B(u) - E\{B(u)\}] = R_B(t,u)$$
 (2.13)

where E denotes the expectation over the ensemble of scenes. Because the detector is nutating, the random process, B(t), will not be stationary. However, if the detector and background do not move with respect to each other, then $R_B(t_1,t_2)$ is doubly periodic:

$$R_B(t_1 + MT, t_2 + NT) = R_B(t_1, t_2)$$
 (2.14)

where M and N are integers.

It can be shown [Ref. 7] that the form of the covariance of a doubly period process is

$$R_{B}(t_{1},t_{3}) = \sum_{m,n}^{\infty} \beta_{mr_{1}} e^{jm\omega_{0}t_{1}} e^{-jn\omega_{0}t_{2}}. \qquad (2.15)$$

Because the nutation is periodic, over the ensemble of scenes, there is little significance in the starting time t. Therefore, considering t as a random variable that is uniform over the nutation period, one can "average it out" of the covariance function as

$$R_{B}(\tau) = \frac{1}{T} \int_{0}^{T} R_{B}(t+\tau,t) dt$$
 (2.16)

Substituting (2.16) in (2.15),

$$R_{B}(\tau) = \sum_{m,n}^{\infty} \beta_{mn} e^{jm\omega_{0}} \frac{1}{T} \int_{0}^{T} e^{j(m-n)\omega_{0}t} dt. \qquad (2.17)$$

Using the relation

$$\delta uv = \frac{1}{T} \int_{0}^{T} e^{j(u-v)\omega_{0}t} dt$$
, (2.18)

the correlation function becomes

$$R_{B}(\tau) = \sum_{n=0}^{\infty} \beta_{n} e^{jn\omega_{0}\tau}, \qquad (2.19)$$

Equation (2.19) shows that B(t) can be considered a wide sense stationary process and that it is periodic in the mean square sense; i.e.,

$$R_{\rm B}(\tau) = R_{\rm B}(T+\tau)$$
 for any τ (2.20)

The power spectral density is the temporal Fourier transform of $R_{\rm B}(\tau)$. Taking the transform one has

$$S_{B}(\omega) = 2\pi \sum_{n=0}^{\infty} \beta_{n} \delta(\omega - n\omega_{0}). \qquad (2.21)$$

The coefficients, \mathbf{S}_{n} , can be related to the optical system parameters by

$$\beta_{n} = \int |\tau(k)|^{2} |F_{0}(k)|^{2} |W_{B}(k)| J_{n}^{2} (2\pi\rho\sqrt{k_{x}^{2} + k_{y}^{2}}) d^{2}k (2.22)$$

where $F_0(k)$ is the transform of the point spread function, and $W_B(k)$ is the Wiener spectrum, the transform of the background correlation function. This derivation was first made by Samuelsson [Ref. 10] in an internal Swedish report then again by Yoshitani [Ref. 6]. Yoshanti's derivation has been included as Appendix B.

Equation (2.19) implies that the background, B(t), can be represented by a Fourier series with uncorrelated coefficients [Ref. 7]. Thus,

$$B(t) = \sum_{n=0}^{\infty} b_{n} e^{jn\omega_{0}t}$$
 (2.23)

where the convergence is in the mean square sense.

The coefficients, b_n , satisfy

$$E(b_n) = \begin{cases} E(B(t)) & n = 0 \\ 0 & n \neq 0 \end{cases} . \qquad (2.24)$$

Moreover, they are uncorrelated

$$E(b_m b_n) = \begin{cases} \beta_n & m = n \\ 0 & m \neq n \end{cases}$$
 (2.25)

C. WHITE NOISE COMPONENT

In an actual system there will be a certain amount of internal noise generated in the detector and preamplifier which is additive to any background noise or signal. In the frequencies of interest the primary source of this noise is thermal agitation. This type of noise, called Johnson noise, is assumed to be white and Gaussian.

Let W(t) be Gaussian white noise with power spectral density $\frac{N}{2}$ (V^2/H_z). The power spectrum at the output of the preamp with bandwidth B is shown in Figure 1.3.

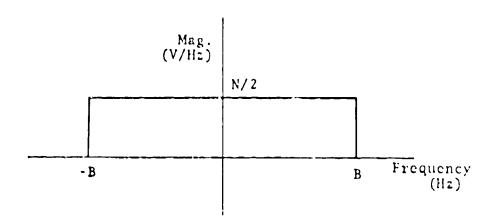


Figure 1.3. Power Spectrum of White Noise.

The RMS voltage measured at the output of the preamp is

$$V^2 \text{rms} = \frac{N}{2} \cdot 2B = NB (V^2)$$
 (2.26)

III. TEMPORAL PROCESSING

The output voltage from the detector is temporally processed to determine the presence of a target. The type of processor is determined by applying statistical detection theory to the known behavior of the nutating system and various realistic assumptions about the system and the target. An optimum processor in the sense of small signal amplitudes, called a threshold detector, is developed in this section.

Harger [Ref. 3] derived the detection theory for the known signal case in the presence of additive background and white noise components. In IR detection, the target amplitude and position are usually not known completely. The threshold detector is an extension of Harger's work which accounts for the unknown amplitude and position.

The form of the optimal detector under the Neyman-Pearson criterion will be derived for multiple observations. It is assumed that the image plane and object plane remain parallel, id stationary with respect to each other during the observations.

The detection of the target under the Neyman-Pearson criterion becomes a problem of hypothesis testing. The task of the detector is one of choosing between two hypotheses, ${\rm H}_0$ that only noise is present and ${\rm H}_1$ that in addition to the noise there is a target present. The design of the detector is one that permits the correct choice of hypothesis ${\rm H}_1$ (a detection) with maximum probability of detection, Qd, in a fixed probability, ${\rm Q}_{\rm FA}$, of choosing

 ${\rm H_1}$ when ${\rm H_0}$ is true (a false alarm). The structure of the Neyman-Fearson criterion requires forming a likelihood ratio, ${\rm A_{10}}$, which is compared to a known threshold, ${\rm V_t}$, specified for a given false alarm probability. If the likelihood ratio exceeds the threshold, hypothesis ${\rm H_1}$ is assumed true and if not, hypothesis ${\rm H_0}$ is assumed true. The logic process may be written

$$\Lambda_{10} \stackrel{H_1}{\underset{H_0}{\stackrel{>}{\triangleright}}} V_{\mathfrak{t}} \tag{3.1}$$

where the threshold, V_{t} , is prived from the expression

$$Q_{FA} = \int_{t}^{\infty} p(Z|H_1) dZ$$
 (3.2)

and Z is a sufficient statistic of the received data.

A. DETECTION OF A KNOWN SIGNAL

The data for processing is a set of M functions of time $\{Z_n(t), n=1,...,M\}$ where each function is the output of the preamp during one nutation (0,T). T is the nutation period. The output may be represented as

$$Z_n(t) = S_n(t) + B(t) + N_n(t); 0 \le t \le T, n=1,...,M$$
 (3.3)

 $S_n(t)$ represents the output due to a target. The component, B(t), represents an additive background and is a sample function of a random field (B). The relative size of the target in the field of view is assumed sufficiently small that the background is additive while B(t) is assumed to be

the sample function during each of the M nutations observed. $N_n(t)$ represents the additive noise component due to the internal noise of the detector-preamp combination. It is a sample function of a white Gaussian random process (N_n) of zero mean and power spectral density $N/2(V^2/H_7)$.

The two hypotheses for the likelihood ratio may be written as

$$H_1: \{Z_n(t) = S_n(t) + B(t) + N_n(t); n=1,...,M\}$$
(3.4)

$$H_0: \{Z_n(t) = B(t) + N_n(t); n=1,...,M$$

Denote the likelihood ratio as $\Lambda_{10}(\{z_n\})$ where the likelihood ratio is defined as

$$\Lambda_{10}(\{z_n\}) = \frac{p(\{z_n\}|H_1)}{p(\{z_n\}|H_0)}$$
 (3.5)

The likelihood ratio is complicated by the fact \mathbf{Z}_n and \mathbf{Z}_m are not independent because of a common noise component $\mathbf{B}(t)$. To simplify the derivation Harger introduced an auxillary hypothesis

$$H_2: \{Z_n = N_n ; n=1,...,M\}.$$
 (3.6)

Then the likelihood ratio, Λ_{10} , can be computed using the chain rule for likelihood ratios:

$$\Lambda_{10} = \frac{p(\{z_n\}|H_1)}{p(\{z_n\}|H_0)} = \frac{p(\{z_n\}|H_1)/p(\{z_n\}|H_2)}{p(\{z_n\}|H_0)/p(\{z_n\}|H_2)}$$
$$= \Lambda_{12}/\Lambda_{02}.$$

If the sample functions for the background, B, were non-random, $Z_i(t)$ and $Z_j(t)$, $i \neq j$, would be mutually independent because of white noise. Accordingly, if B were fixed, one can write the conditional likelihood for the n^{th} observation Λ_{12}^n ($\{Z_n\}|B$). Moreover, the likelihood ratio for M observations would be the product of the likelihood ratios for each single observation,

$$\Lambda_{12} (\{Z_n\} | B) = \prod_{n=1}^{M} \Lambda_{12}^n (Z_n | B).$$

The conditioning is then removed by averaging over the ensemble of backgrounds

$$\Lambda_{12}(\{z_n\}) = E_B[\Lambda_{12}(\{z_n\}|B)].$$

Since the noise is white Gaussian, the conditional likelihood ratio is that for detecting a known signal $S_n^+B_n^-$ in white Gaussian noise and has the well known form ([Ref. 15], page 253)

$$\Lambda_{12}^{n} (\{Z_{n}\} | B) = \exp \left\{ \frac{2}{N} \sum_{n=1}^{M} \int_{0}^{T} Z_{n}(t) (S_{n}(t) + B(t)) dt - \frac{1}{N} \sum_{n=1}^{M} \int_{0}^{T} (S_{n}(t) + B(t))^{2} dt \right\}$$

which may be rewritten as

$$\Lambda_{12}^{n}(\{Z_{n}\}|B) = \exp\left\{\frac{2}{N}\sum_{n=1}^{M}\int_{0}^{T}Z_{n}(t) S_{n}(t) dt\right\} - \frac{1}{N}\sum_{n=1}^{M}\int_{0}^{T}S_{n}^{2}(t) dt \exp\alpha(b)$$
 (3.14)

where

$$\alpha(b) = \frac{2}{N} \sum_{n=1}^{M} \int_{0}^{T} (Z_{n}(t) + S_{n}(t)) B(t) dt - \frac{M}{N} \int_{0}^{T} B^{2}(t) dt.$$
 (3.15)

To rind $\Lambda_{10}(\{Z_n\})$, one must evaluate $E_B[\Lambda^n_{\ 12}(\{Z_n\}|B)]$ where

$$\begin{split} E_{B}[\Lambda^{n}_{12}(\{2_{n}\}|P)] &= \exp \{\{\frac{2}{N}\sum_{n=1}^{M}\int_{0}^{T}Z_{n}(t)|S_{n}(t)|dt\} \\ &- \frac{1}{N}\sum_{n=1}^{M}\int_{0}^{T}|S_{n}^{2}(t)dt\} \cdot E_{B}\{\exp \alpha(b)\}. \end{split}$$
 (3.16)

Assuming the background is Gaussian the expectation can be evaluated by representing the background in a Karhunen-Loeve (K-L) expansion

$$B(t) = \sum_{k} b_{k} \phi_{k}(t) , 0 \le t \le T$$
 (3.17)

where $\phi_{\bf k}(t)$ is the eigenf nction corresponding to eigenvalue $\lambda_{\bf k}$ of the homogeneous integral equation

$$\lambda_{k} \phi_{k}(t) = \int_{0}^{T} R_{B}(t, u) \phi_{k}(u) du, \quad 0 \leq t \leq T. \quad (3.18)$$

 $R_{\mbox{\footnotesize{B}}}(\mbox{t,u})$ is the covariance function of background

$$R_B(t,u) = E\{[B(t) - m_B(t)][B(u) - m_B(u)]\}$$
 (3.19)

where $m_B(t)$ is the mean of the background.

Then $E_B(\exp\alpha(b))$ can be calculated in a straightforward manner. After the lengthy calculation shown in Appendix C

$$\ln E_B \{\exp \alpha(b)\} = -\frac{1}{2} \sum_{k} \ln (1 + \frac{M}{N_0} \lambda_k)$$

$$+ \frac{2}{N} \int_{0}^{T} dt \sum_{n=1}^{M} (Z_{n}(t) - S_{n}(t)) \int_{0}^{T} du \sum_{m=1}^{M} (Z_{m}(u) - S_{m}(u)) q(t,u)$$

$$-\frac{1}{2}\int_{0}^{T} dt \left\{ m_{B}(t) + \frac{2}{M}\sum_{n=1}^{M} (Z_{n}(t) - S_{n}(t)) \right\} h(t)$$
 (3.20)

where q(t,u) satisfies the integral equation

$$\frac{N}{2M} q(t,u) + \int_{0}^{T} R_{B}(t,\tau)q(\tau,u)d\tau = R_{B}(t,u) \quad 0 \le u, t \le T \quad (3.21)$$

and h(t) satisfies

$$\frac{N}{2M} h(t) + \int_{0}^{T} R_{B}(t,\tau)h(\tau)d\tau = m_{B}(t) \quad 0 \le t \le T. \quad (3.22)$$

q(t,u) is the impulse response of the filter that gives the minimum mean square error estimate \hat{E} of the background when the mean is zero [Ref. 15]. The additive white noise is reduced in amplitude by 1/M. Using the K-L representation, one can also write

$$q(t,u) = \sum_{k=1}^{\infty} \frac{\frac{N}{2M} \lambda_k}{\frac{N}{2M} + \lambda_k} \phi_k(t) \phi_k^*(u) \quad 0 \le u, t \le T$$
 (3.23)

Using equations (2.9), (3.10), (3.14) and (3.20) and

observing that

$$\Lambda_{02}(\{Z_n\}) = \Lambda_{12}(\{Z_n\}) | S_n = 0$$
 (3.24)

it can be shown (see Appendix C) that the likelihood function for hypotheses ${\rm H}_1$ and ${\rm H}_0$ is

$$\Lambda_{10}(\{Z_n\}) = \exp \{\sum_{n=1}^{M} \sum_{0}^{T} Z_n(t)q_n(t)dt + K'\}$$
 (3.25)

where

$$q_n(t) = \frac{2}{N} [S_n(t) - \frac{1}{M} \int_{0}^{T} \sum_{\ell=1}^{M} S_{\ell}(u)q(t,u)du]$$
 (3.25a)

and

$$K' = -\frac{1}{N} \sum_{n=1}^{M} \int_{0}^{T} S_{n}^{2}(t) dt + \frac{1}{NM} \int_{0}^{T} \sum_{n=1}^{M} \sum_{n=1}^{M} S_{n}(t) S_{2}(u) q(t, u) dt du$$

$$-\frac{2}{m} \sum_{n=1}^{M} \int_{0}^{T} S_{n}(t) h(t) dt \qquad (3.25b)$$

The logarithm is a monotonic function. Thus it offers no loss of information but greatly simplifies the calculations. In the remaining work

$$\Lambda_{10}^{\prime} (\{Z_n\}) = \ln \Lambda_{10} (\{Z_n\})$$
 (3.26)

Rewriting equation (3.25),

$$\Lambda'_{10}(\{z_n\}) = L(\{z_n\}) + K'$$
 (3.27)

where

$$L(Z_n) = \sum_{n=1}^{M} \int_{0}^{T} Z_n(t)q_n(t)dt.$$

K' does not depend on the input data so it may be included in the threshold. The hypotheses test now becomes

$$L(Z_n) \stackrel{H_1}{\stackrel{<}{\sim}} Threshold . \qquad (3.28)$$

It is also possible, as Harger has shown, to derive tests similar to (3.28) without making the Gaussian background assumption. In particular, if the background fluctuates many times over (0, T) such that $\alpha(b)$ of (3.15) is approximately Gaussian, then the test is

$$\begin{array}{ccc}
M & T \\
\Sigma & \int Z_{n}(t)g_{n}(t)dt & \\
n=1 & 0 & H_{0}
\end{array}$$
Threshold
(3.29)

where

$$g_{n}(t) = \frac{1}{N} [S_{n}(t) - \frac{2M}{N} \int_{0}^{T} R_{B}(t,u) [\frac{1}{M} \sum_{k=1}^{M} S_{k}(u)]du].$$

Since by Mercer's theorem

$$R_{B}(t,u) = \sum_{k}^{\infty} \lambda_{k} \phi_{k}(t) \phi_{k}^{*}(u), \quad 0 \leq t, u \leq T$$
 (3.30)

on comparing $\mathbf{q}_{\mathbf{n}}$, $\mathbf{g}_{\mathbf{n}}$ and (3.21), it is seen that the above model approaches the Gaussian model if

$$\frac{N}{2M} \gg \max_{k} \lambda_{k}$$
.

The above condition, which implies a weak background, is sufficient but by no means necessary. The two models approximate each other whenever the second term of \mathbf{q}_n and \mathbf{g}_n are nearly identical. The point is that while the Gaussian background assumption may not fit the reality, a processor based on the assumption is probably not far from being the optimum. Also, the Gaussian assumption allows one to compute the processor performance.

B. DETECTION OF SIGNAL WITH UNKNOWN PARAMETERS

The likelihood ratio derived in section A is applicable only if the signal is known completely. Such is seldom the case in IR target detection and in general the signal is a function of some unknown parameters. Typically, the amplitude, a, of the target and its position \vec{r}_0 in the image plane are the unknown parameters, and the form of the signal can be rewritten to include these parameters:

$$S_n = S_n(t;a, \dot{\tau}_0)$$

Furthermore, by assuming that only point targets are of interest and including the assumption of fixed image and object planes over the number of nutations of interest, the shape of the waveform will be the same during each nutation. However, the amplitude is allowed to vary between successive nutations.

The signal waveform may now be written as

$$S_n(t; a, \vec{r}_o) = a_n f(t, \vec{r}_o)$$
 (3.31)

where a_n is the signal amplitude and $f(t, \vec{r}_0)$ is the signal shape for a target at \vec{r}_0 .

Rewriting hypothesis H_1 under these assumptions gives

$$H_1': \{Z_n(t) = a_n f(t, \vec{\tau}_0) + B(t) + N_n(t); n=1,...,M$$

$$0 \le t \le T\}$$
(3.32)

The likelihood ratio Λ_{10} becomes the conditional likelihood ratio $\Lambda_{10}(\{Z_n\}|a,\vec{r}_o)$ obtained by replacing $S_n(t)$ by $a_nf(t,\vec{r}_o)$ everywhere in (3.25).

Assuming a priori statistical knowledge of a and \dot{r}_0 , and their independence, the likelihood ratio may be averaged over the respective density functions to produce an unconditional likelihood ratio

$$\Lambda_{10}(\{Z_n\}) = \iint \Lambda_{10}(\{Z_n\}|a, \vec{r}_o)p(a)p(\vec{r}_o)dad\vec{r}_o.$$
 (3.33)

Instead of averaging as above, one could estimate a and $\dot{\vec{r}}_{o}$ by the maximum likelihood procedure and substitute the

estimated values $\hat{\mathbf{a}}$ and $\hat{\mathbf{r}}_0$ in the conditional likelihood ratio and try to maximize the generalized likelihood ratio $\Lambda_{10}(\{\mathbf{Z}_n\}|\hat{\hat{\mathbf{a}}},\hat{\hat{\mathbf{r}}}_0)$. However, this procedure is not followed here because the maximum likelihood estimates are difficult to obtain.

Since the main interest is in detection of the weak target, the small signal case is of primary interest. Expanding $\Lambda_{10}(\{Z_n\}|a,\dot{\tau}_0)$ in powers of signal amplitudes

$$\Lambda_{10}(\{Z_{n}\}|\underline{a}, \overrightarrow{r}_{0}) = 1 + \sum_{n=1}^{M} a_{n} \frac{\partial}{\partial a_{n}} \Lambda_{10}(\{Z_{n}\}|\underline{a}, \overrightarrow{r}_{0})|_{\underline{a}=0}$$

$$+ \frac{1}{2} \sum_{n=1}^{M} \sum_{k=1}^{M} a_{n} a_{k} \frac{\partial^{2}}{\partial a_{n} \partial a_{k}} \Lambda_{10}(\{Z_{n}\}|\underline{a}, \overrightarrow{r}_{0})|_{\underline{a}=0}$$
(3.34)

*****

Integrating (3,34) as in (3.33) term by term and retaining only the first non-zero term that depends on the data provides a statistic which is optimum for weak signals and is called a threshold detector [Ref. 4].

The evaluation is made easier by the fact

$$\frac{\partial}{\partial a_{k}} \Lambda_{10}(\{Z_{n}\}|a, r_{0})| = \frac{\partial}{\partial a_{k}} \ln \Lambda_{10}(\{Z_{n}\}|a, r_{0})| = 0$$
(3.35)

which results from

$$\Lambda(\{z_n\}|_{\infty}^{0}, \dot{r}_0) = 1$$
.

The first non-zero term becomes

$$\frac{\partial}{\partial a_{k}} \ln \Lambda_{10}(\{Z_{n}\}|\underline{a}, \dot{r}_{0})|_{\underline{a}=0} = \frac{2}{N} \int_{0}^{T} Z_{n}(t) f(t, \dot{r}_{0}) dt$$

$$- \frac{2}{N} \int_{0}^{T} Z_{n}(t) dt \int_{0}^{T} q(t, u) f(u, \dot{r}_{0}) du$$

$$- \int_{0}^{T} f(t, \dot{r}_{0}) h(t) dt \qquad (3.36)$$

Performing the integration over \dot{r}_0 ,

$$\Lambda_{10}(\{Z_{n}\}) = 1 + \frac{2}{N} \sum_{n=1}^{M} \bar{a}_{n} \int_{0}^{T} Z_{n}(t) [\bar{f}(t) - \int_{0}^{T} q(t, u) \bar{f}(u) du]$$

$$-\int_{0}^{T} \bar{f}(t) h(t) dt \qquad (3.37)$$

where the bar denotes the averaged quantities. The first and last terms are independent of the data and may be included in the threshold value. Moreover, if the amplitudes are identically distributed, as is generally the case, the threshold detector performs the test

$$W = \sum_{n=1}^{M} \int_{0}^{T} Z_{n}(t) \dot{q}(t) dt + W_{t}$$

$$H_{0}$$
(3.38)

where W_{t} is the threshold and

$$\bar{q}(t) = \frac{2}{N} \left[\bar{f}(t) - \int_{0}^{T} q(t, u) \bar{f}(u) du \right]. \qquad (3.59)$$

Thus the threshold detector can be mechanized either as a correlator or matched filter followed by an accumulator.

The filter function $\tilde{q}(t)$ may be represented in a series expansion in terms of the eigenfunctions of $R_B(t,u)$ and the coefficients of the signal expansion. Expanding $\tilde{f}(t)$ in terms of the eigenfunctions:

$$\bar{f}(t) = \sum_{k}^{\infty} \bar{f}_{k} \phi_{k}(t) . \qquad (3.40)$$

Substituting (3.40) in (3.39) and using (3.23)

$$\bar{q}(t) = \frac{N}{2M} \sum_{k}^{\infty} \frac{\bar{f}_{k}}{\frac{N}{2M} + \lambda_{k}} \phi_{k}(t) . \qquad (3.41)$$

C. EVALUATION OF PERFORMANCE

The Gaussian assumption was made in deriving the threshold statistic; therefore, the threshold statistic itself is Gaussian. This means that the probability of detection, Qd, and probability of false alarm, Q_{FA} , are uniquely determined by second-order statistics, e.g., the means and variances of W under hypotheses H_0 and H_1 .

Recall that the threshold statistic is

$$W = \sum_{n=1}^{M} \sum_{0}^{T} Z_{n}(t) \bar{q}(t) dt. \qquad (3.42)$$

The means and variances of W are:

$$E(W|H_0) = \sum_{n=1}^{M} \int_{0}^{T} \tilde{q}(t) E(B(t) + N_n(t)) dt$$
$$= \sum_{n=1}^{M} \int_{0}^{T} \tilde{q}(t) m_B(t) dt$$

$$E(W|H_1) = \sum_{n=1}^{M} \int_{0}^{T} \tilde{q}(t) E[a_n f(t,r_0) + B(t) + N_m(t)] dt$$

$$= E(W|H_0) + \sum_{n=1}^{M} \int_{0}^{T} \tilde{f}(t) \tilde{q}(t) dt$$

 $VAR (W|H_0) = VAR (W|H_1) = VAR$

$$= \frac{NM}{2} \int_{0}^{T} \bar{q}^{2}(t) dt + M^{2} \int_{0}^{T} \bar{q}(t) \int_{0}^{T} \bar{q}(u) R_{B}(t,u) dudt$$
(3.43)

Letting $p_w(\cdot 1H_i)$ be the Gaussian density function of W under H_i which is specified by the equations above, then

$$Q_{FA} = \int_{W_{t}}^{\infty} p_{w}(x1H_{0}) dx \qquad (3.44)$$

and

$$Q_{d} = \int_{W_{+}}^{\infty} p_{W}(x1H_{1}) dx$$
 (3.45)

where W_{+} is the threshold voltage.

Equations (3.43) and (3.44) can be solved to yield

$$Q_{FA} = \phi_C(x) \tag{3.46}$$

and

$$Q_{d} = \phi_{C}(x-d) \tag{3.47}$$

where-

$$x = \frac{W_t - E[W|H_0]}{\sqrt{VAR}},$$

d is the equivalent signal-to-noise ratio

$$d = \frac{E(W|H_1]}{\sqrt{VAR}}$$

$$\frac{\sum_{n=1}^{M} a_{n} \int_{0}^{T_{-}} f(t) \tilde{q}(t) dt}{\sqrt{\frac{NM}{2} \int_{0}^{T} \tilde{q}^{2}(t) dt + M^{2} \int_{0}^{T} \tilde{q}(t) \int_{0}^{T} \tilde{q}(u) R_{B}(t,u) dudt}}$$
(3.48)

and $\phi_{\boldsymbol{c}}(g)$ is the complementary error function of the kind

$$\phi_{\mathbf{d}}(g) = \frac{1}{\sqrt{2\pi}} \int_{g}^{\infty} e^{-v^{2/2}} dv$$
 (3.49)

By setting a desired probability of false alarm, x may be determined from equation (3.46). The probability of detection is then determined by solving for d and using equation (3.47).

By substituting (3.41) and (3.42), and using (3.23), the numerator of d is

$$\sum_{n=1}^{M} a_n \int_{0}^{T} f(t)q(t)dt = \sum_{n=1}^{M} \frac{N}{2M} \sum_{k} \frac{|\tilde{f}_{k}^{*}|^2}{N + \lambda_k}$$
(3.50a)

while the denominator of d (see Appendix D) is

$$\sqrt{VAR} = \frac{N}{2M} \left[\sum_{k=1}^{\infty} \frac{|\hat{f}_{k}|^{2}}{\frac{N}{2M} + \lambda_{k}} \right]^{\frac{1}{2}}$$
(3.50b)

Therefore the effective signal-to-noise ratio is

$$d = \frac{1}{M} \sum_{n=1}^{M} \tilde{a}_n \sqrt{\frac{2M}{N}} \left[\sum_{k=1}^{\infty} \frac{|\tilde{f}_k|^2}{1 + \frac{2M}{N} \lambda_k} \right]^{l_2}. \tag{3.51}$$

Knowledge of d along with (3.46) and (3.47) completely specifies the performance of the threshold detector.

The equivalent signal-to-noise ratio for a target whose amplitude and position are known is easily shown to be

$$d' = \frac{1}{M} \sum_{n=1}^{M} a_n \sqrt{\frac{2M}{N}} \left[\sum_{k=1}^{\infty} |f_k(\vec{r}_0)|^2 \frac{1}{1 + \frac{2M}{N} \lambda_k} \right]^{k}$$
 (3. 2)

By letting the background be zero it will now be shown that d' reduces to the equivalent signal-to-noise ratio for a known signal in white Gaussian noise which is well known [Ref. 4].

The energy dissipated during one observation in a resistor of 1 ohm if the signal $a_n f(t, \vec{r}_0)$ is the voltage across that resistor is

$$\underline{E}_{n} = a_{n}^{2} \int_{0}^{T} \sum_{k}^{\infty} f_{k}(\vec{r}_{o}) \phi_{k}(t) \sum_{k}^{\infty} f_{k}^{*} (\vec{r}_{o}) \phi_{k}^{*}(t) dt$$

$$= a_{n}^{2} \sum_{k}^{\infty} |f_{k}(\vec{r}_{o})|^{2}. \qquad (3.53)$$

The voltage resulting from M observations then is

$$\sqrt{E} = \frac{1}{M} \sum_{n=1}^{M} \sqrt{E_n} - \frac{1}{M} \sum_{n=1}^{M} a_n \left[\sum_{k=1}^{\infty} |f_k(\vec{r}_0)|^2 \right]^{\frac{1}{2}}. \quad (3.54)$$

Substituting (3.54) in (3.52) with $\lambda_k = 0$,

$$d' = \sqrt{E} \quad \sqrt{\frac{2M}{N}} = \sqrt{\frac{2E}{N}} \sqrt{M} ,$$

which shows the familiar result that detectability depends only upon the signal energy and the signal-to-noise ratio increases as the square root of the number of observations.

IV. THRESHOLD DETECTOR FOR NUTATING SYSTEM

In an infrared detection system the output voltage to be processed is obtained from a detector-preamplifier combination. The irradiance on the detector in W/cm^2 is converted to volts with a linear scale factor. The voltage output can be written as

$$v(t, \vec{r}_0) = KH(t, \vec{r}_0)$$
 (4.1)

where $H(t, \vec{r}_0)$ is the irradiance on the detector from a target at position \vec{r}_0 , and K is a scale factor in $(V/W/cm^2)$. Also the covariance function may be written as

$$R_{VB}(\tau) = E[KB(t+\tau) KB(t)] = K^2 R_B(\tau),$$
 (4.2)

where B(t) is the irradiance on the detector due to a particular background scene. The output voltage due to this scene is

$$v_{R} = KB(t) \tag{4.3}$$

The problem now is to relate the detector-preamplifier output to the equations derived for the threshold detector.

It was shown in Part III that an optimum processor for weak signals was a threshold detector that performed a certain linear operation on the data. The linear filter was specified in terms of the eigenvalues and eigenfunctions of the integral equation (3.18), where the kernal $R_{\rm B}({\rm t,u})$ is the covariance function for the background.

In Part II, the covariance function was noted to be periodic. Rewriting (4.2) in terms of the output voltage

$$R_{VB}(t,u) = \kappa^2 \sum_{k}^{\infty} \beta_k e^{jk\omega_0(t-u)}$$
 (4.4)

Substituting (4.4) into integral equation (3.18), the solution for the eigenvalues and eigenfunctions is trivial with

$$\phi_k(t) = \frac{e^{jk\omega_0 t}}{\sqrt{T}}$$

and

$$\lambda_{k} = \chi^{2} \beta_{k} T \tag{4.5}$$

where T is the nutation period.

The signal coefficients were also noted to be periodic and could be expanded in a Fourier series

$$v(t) = K \sum_{k}^{\infty} H_{k}(\hat{r}_{o}) e^{jk\omega_{o}t}.$$
 (4.)

Recognizing the similarity between equations (4.6) and (3.1), the signal coefficients for a known signal which is expanded using the basis set above become

$$f_k(\vec{r}_0) = KH_k(\vec{r}_0) \sqrt{T}$$
 (4.7)

For the threshold detector, the signal is averaged with respect to an a priori density function for \vec{r}_0 . Averaging can be done with the coefficients to give

$$\hat{\mathbf{f}}_{\mathbf{k}} = \int p(\vec{r}_{\mathbf{o}}) \mathbf{f}_{\mathbf{k}}(\vec{r}_{\mathbf{o}}) d\vec{r}_{\mathbf{o}}$$

$$= K \sqrt{T} \int P_{\mathbf{k}}(\vec{r}_{\mathbf{o}}) p(\vec{r}_{\mathbf{o}}) d\vec{r}_{\mathbf{o}}$$

$$= K \sqrt{T} \tilde{P}_{\mathbf{k}} .$$
(4.8)

Using equation (2.10), the coefficients $\tilde{\mathbf{H}}_k$ can be written as

$$\bar{H}_{k} = j^{n} \int d\vec{r}_{o} p(\vec{r}_{o}) \int N'(k, \vec{r}_{o}) \tau^{*}(k) e^{-jn\phi} J_{n} \left(2\pi\rho \sqrt{k_{x}^{2} + k_{y}^{2}}\right) d^{2}k \qquad (4.9)$$

The transform of the radiance distribution from a target, $N'(k,\vec{r}_0)$ is a function of the target radiance, R(k), the position of the target, \vec{r}_0 , and the point spread function of the optics, $F_0(k)$. Rewriting (4.9),

$$\tilde{H}_{k} = j^{n} \int d^{2}k R(k) \tau^{*}(k) J_{n} \left(2\pi\rho \sqrt{k_{x}^{2} + k_{y}^{2}}\right) e^{-jn\phi}$$

$$\int d\tilde{r}_{o} p(\tilde{r}_{o}) e^{-j2\pi k_{o}^{2}, \tilde{r}_{o}^{2}} \qquad (4.10)$$

If the density function, $p(\vec{r}_0)$, is assumed Gaussian with a standard deviation p_0 , then the averaged coefficients become

$$\tilde{H}_{k} = j^{n} \int R(k) e^{-2\pi p_{0}^{2}(k_{x}^{2} + k_{y}^{2})} \tau^{*}(k) e^{-jn\phi} J_{n} \left(2\pi \rho \sqrt{k_{x}^{2} + k_{y}^{2}}\right) d^{2}k. \quad (4.11)$$

The Gaussian assumption on \vec{r}_0 is reasonable because \vec{r}_0 often represents a pointing error from some designated target position. Then p_0 describes the accuracy of the pointing mechanism. Substituting (4.8) into (3.42), the filter function for the threshold detector becomes

$$\ddot{q}(t) = K \sum_{k=1}^{\infty} \ddot{\Pi}_{k} \sqrt{T} \left(\frac{1}{1 + \frac{2M}{N} K^{2} \beta_{k} T} \right) \frac{e^{j k \omega_{o} t}}{\sqrt{T}}$$
(4.12)

Rewriting the equivalent signal-to-noise ratio in terms of equations (4.8) and (4.12) one has

$$d = \frac{\frac{M}{\sum_{n=1}^{K} \bar{a}_{n}} \sqrt{T} \sqrt{M}}{\sqrt{N/2}} \left[\sum_{k} |\bar{H}_{k}|^{2} \frac{1}{1 + \frac{MTK^{2}\beta_{k}}{N/2}} \right]^{\frac{1}{2}}$$
(4.13)

It will be shown in Part V that $\boldsymbol{\beta}_k$ may be represented as

$$\varepsilon_k = \sigma_B^2 \Omega^2 Q_k$$

where $\sigma_p^{\ 2}$ is the amplitude of the background noise correlation function, Ω^2 is the instantaneous field-of-view of the detector and Q_k is an integral function of the optical and detector parameters and correlation length of the background radiance.

Define the following terms:

Signal-to-noise ratio (SNR) =
$$\frac{K \frac{1}{M} \sum_{n=1}^{M} \tilde{a}_n \sqrt{T}}{\sqrt{N/2}}$$
 (4.14a)

Background-to-noise ratio (BNR) =
$$\frac{\sigma_B \Omega K \sqrt{T}}{\sqrt{N/2}}$$
 (4.14b)

Using (4.14a) and (4.14b), d may be rewritten as

$$d = SNR \sqrt{M} \left[\sum_{k} |\bar{H}_{k}|^{2} \frac{1}{1 + MQ_{k}PNR^{2}} \right]^{\frac{1}{2}}$$
 (4.15)

Noise equivalent irradiance (NEI) may be defined at the detector-preamplifier output to be

$$K \cdot NEI = \sqrt{\frac{N}{2} \cdot 2B_{W}} = \sqrt{NB}_{W}$$
 (4.16)

where $\boldsymbol{B}_{\boldsymbol{w}}$ is the detector-preamplifier bandwidth.

The SNR and BNR may now be written in terms of the NEI as

$$\sqrt{M}$$
 SNR = $\frac{\langle a \rangle \sqrt{NTB}}{NEI}$ = SNR' \sqrt{NTB}

where NTB is number-time-bandwidth product and

$$\sqrt{M}$$
 BNR = $\frac{\sigma_R}{NET}$ \sqrt{NTB} = BNR' \sqrt{NTB} .

Expressing d in terms of NEI one has

$$d = SNR' \sqrt{NTP} \sum_{k}^{\infty} |H_{k}|^{2} \left[\frac{1}{1 + Q_{k} NTB BNR'^{2}} \right]^{\frac{1}{2}}.$$
 (4.17)

V. SOLUTION FOR RECTANGULAR DETECTOR

The equations in Part II for the detector-preamplifier output voltage and the correlation function of the background must be solved to obtain the filter function and specify the detection system performance.

The two general equations obtained in Part II are

$$H_{n}(\vec{r}_{o}) = j^{n} \int R(k) e^{-j\pi k \cdot \vec{r}_{o}} \vec{F}_{o}(k) \tau^{*}(k) e^{-j\pi \phi}$$

$$J_{n}\left(2\pi\rho\sqrt{k_{x}^{2} + k_{y}^{2}}\right) d^{2}k .$$

where $H_n(\vec{r}_0)$ is the nth Fourier series coefficient for the radiation on the spatial filter and

$$\beta_{n} = \int |\tau(\underline{k})|^{2} |F_{o}(\underline{k})|^{2} |W_{B}(\underline{k})|^{2} \int_{0}^{2} d^{2}\underline{k} d^{2}\underline{k}$$

where B_n is the n^{th} coefficient of the background correlation function at the spatial filter.

A rectangular shaped detector was chosen for the calculations because it exhibited the simplest form of the equations. For circular nutation the filter is represented graphically in Figure 5.1.

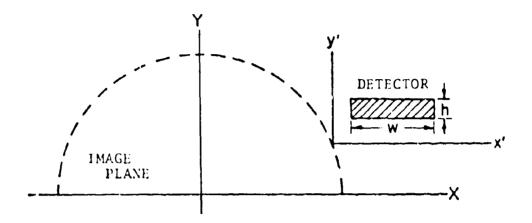


Figure 5.1. Nutation with Rectangular Detector.

If the detector function is given by

$$\tau(x',y') = \begin{cases} 1 & x_0-w/2 \le x' \le x_0+w/2 \\ y_0-h/2 \le y' \le y_0+h/2 \end{cases}$$
0 otherwise

then its transform is

$$\tau(k_{x},k_{y}) = \frac{e^{-j2\pi(k_{x}x_{0}+k_{y}y_{0})}}{\pi^{2}k_{x}k_{y}} \sin(\pi k_{x}w)\sin(\pi k_{y}h)$$

so that

$$|\tau(k)|^2 = \frac{\sin^2(\pi k_x w)}{\pi^2 k_x^2} \frac{\sin^2(\pi k_y h)}{\pi^2 k_y^2}$$
 (5.1)

The point spread function of the optics is assumed to be Gaussian:

$$f_o(\mathbf{r}) = \frac{1}{2\pi\sigma^2} e^{-\frac{|\mathbf{r}|^2}{2\sigma^2}}$$

Its transform is

$$F_{o}(k) = e^{-2\pi^{2}\sigma^{2}} (k_{x}^{2} + k_{y}^{2})$$
 (5.2)

The correlation of the background radiance is assumed to be of the form

$$\phi_{B}(x,y) = \sigma_{B}^{2} e^{-\alpha |x|} e^{-\beta |y|}$$
 (5.3a)

where α^{-1} and β^{-1} are the correlation lengths in the x- and y- directions, respectively. The Wiener spectrum is the Fourier transform of ϕ_R :

$$W_{B}(k) = \sigma_{B}^{2} \frac{2\alpha}{\alpha^{2} + (2\pi k_{x})^{2}} \frac{2\beta}{\beta^{2} + (2\pi k_{y})^{2}}$$
 (5.3b)

Substituting (5.1), (5.2) and (5.3b) into (2.10), the radiance from a point target becomes

$$H_{n}(\vec{r}_{o}) = j^{n} \int e^{-j2\pi(k_{x}r_{x}+k_{y}r_{y})} e^{-2\pi^{2}\sigma^{2}(k_{x}^{2}+k_{y}^{2})} e^{-j2\pi(k_{x}r_{x}+k_{y}r_{y})} e^{-j2\pi^{2}\sigma^{2}(k_{x}^{2}+k_{y}^{2})} e^{-jn\pi(k_{x}r_{x}+k_{y}r_{y})} e^{-jn\pi(k_{x}r_{x}+k_$$

Letting

$$u_{x} = \sqrt{2}\pi\sigma k_{x}$$

$$u_{y} = \sqrt{2}\pi\sigma k_{y}$$

equation (5.4) becomes

$$H_{n}(\dot{r}_{o}) = \frac{j^{n}}{\pi^{2}} \int e^{-(u_{x}^{2} + u_{y}^{2})} e^{-j\frac{\sqrt{2}}{\sigma} \left[(r_{x} - x_{o})u_{x} + (r_{y} - y_{o})u_{y} \right]} e^{-jn - tan^{-1}(u_{y}/u_{x})} J_{n} \left(\frac{\sqrt{2}\sigma}{\sigma} - \sqrt{u_{x}^{2} + u_{y}^{2}} \right) du_{x} du_{y}$$
 (5.5)

If \vec{r}_o is random, the radiance from a target may be averaged with respect to the density function $p(\vec{r}_o)$ as shown in Part IV. In particular, if \vec{r}_o is Gaussian distributed, with standard deviation p_o the averaged coefficients become

$$\bar{H}_{n} = j^{n} \int e^{-2\pi^{2}\gamma^{2}(k_{x}^{2}+k_{y}^{2})} e^{j2\pi(k_{x}x_{0}+k_{y}y_{0})}$$

$$= \frac{\sin^{n}k_{x}w}{\pi k_{x}} \frac{\sin^{n}k_{y}h}{\pi k_{y}} e^{-jn \tan^{-1}(k_{y}/k_{x})}$$

$$J_{n} \left(2\pi\rho\sqrt{k_{x}^{2}+k_{y}^{2}}\right) dk_{x}dk_{y}$$
(5.6)

where

$$y^2 = p_0^2 + \sigma^2$$
.

Letting

$$u_x' = \sqrt{2}\pi\gamma k_x$$

$$u_y' = \sqrt{2}\pi \gamma k_y$$

equation (5.6) becomes

$$\bar{H}_{n} = \frac{i}{\pi^{2}} \int e^{-(u_{x}^{'2} - u_{n}^{'2})} e^{j\frac{\sqrt{2}}{\gamma}(u_{x}^{'} x_{o}^{+} u_{y}^{'} y_{o}^{-})} e^{-jn \tan^{-1}(u_{x}^{'} / u_{y}^{'})}$$

$$\cdot J_{n} \left(\frac{\sqrt{2}p}{\gamma} \sqrt{u_{x}^{'2} + u_{y}^{'2}} \right) du_{x}^{'} du_{y}^{'}$$
(5.7)

And finally substituting equations (5.1), (5.2) and (5.3b) into equation (2.22), the correlation coefficients become

$$\beta_{n} = \int e^{-4\pi^{2}\sigma^{2}(k_{x}^{2}+k_{y}^{2})} \frac{\sin^{2}(\pi k_{x}w)}{(\pi k_{x})^{2}} \frac{\sin^{2}(\pi k_{y}h)}{(\pi k_{y})^{2}}$$

$$\sigma_{B}^{2} \frac{2\alpha}{\alpha^{2} + (2\pi k_{x})^{2}} \frac{2\beta}{\beta^{2} + (2\pi k_{y})^{2}} J_{n}^{2} \left(2\pi\rho\sqrt{k_{x}^{2}+k_{y}^{2}}\right) dk_{x} dk_{y}$$
(5.8)

Letting

$$\mathbf{u}_{\mathbf{X}}^{"} = 2\pi\rho \mathbf{k}_{\mathbf{X}}$$

$$\mathbf{u}_{\mathbf{V}}^{"} = 2\pi\rho \mathbf{k}_{\mathbf{V}}$$

equation (5.8) becomes

$$\beta_{n} = \frac{\sigma_{B}^{2} \Omega^{2} \alpha \beta \sigma^{2}}{\pi^{2}} \int e^{-\frac{i_{x}^{2} - u_{y}^{2}}{x}} \left(\frac{\sin \frac{w}{\sqrt{2}\sigma} u_{x}^{"}}{\frac{w}{\sqrt{2}\sigma} u_{x}^{"}} \right)^{2} \left(\frac{\sin \frac{h}{\sqrt{2}\sigma} u_{y}^{"}}{\frac{h}{\sqrt{2}\sigma} u_{y}^{"}} \right)^{2}$$

$$= \frac{1}{(\alpha\sigma)^{2} + u_{x}^{"2}} \frac{1}{(\beta\sigma)^{2} + u_{y}^{"2}} J_{n}^{2} \left(\frac{\rho}{\sigma} \sqrt{u_{x}^{2} + u_{y}^{2}} \right) du_{x}^{"} du_{y}^{"}$$
 (5.9)

where

$$\Omega^2 = w^2 h^2. \tag{5.9}$$

The integrals in (5.5), (5.7) and (5.9) were difficult to evaluate. Although the components of the integrand are relatively simple functions, the argument of the Bessel function prohibited separating the integral into two one-dimensional integrals. No closed form solutions were found so the integrals were evaluated numerically. The numerical method used is based on work by Pierce [9] who applied Gaussian quadrature formulas to two dimensional integration by integrating over a planar annulus in the (x,y) plane.

Gaussian quadrature formulas are a means of evaluating the integral by summing weighted values of the integrand at specific points. For the one dimensional case,

where g(x) is the integrand to be integrated and w(x) is a weight function for which the specific integration was derived. For well behaved functions the integral may be evaluated as

$$\int_{\mathbf{g}}^{\mathbf{b}} \mathbf{g}(\mathbf{x}) d\mathbf{x} = \int_{\mathbf{i}}^{\mathbf{m}} \mathbf{A}_{\mathbf{i}} f(\mathbf{x}_{\mathbf{i}}) + \text{error}$$
(5.11)

where the x_i 's are the m zeros of the mth polynomial

$$P_{m}(x) = \prod_{i=1}^{m} (x-x_{i})$$

of the set of polynomials mutually orthogonal on the interval [a,b] with respect to the weight function w(x). The weights are given by

$$A_{i} = \int_{a}^{b} w(x) L_{i}(x) dx \qquad (5.13)$$

where $L_i(x) = P_m(x)/[(x-x_i)P_m(x_i)]$ is the Lagrange interpolation coefficient.

For a weight function w(x) and interval [a,b] the polynomial $P_m(x)$ and its zeros x_i and weight factors A_i need be computed only once. For a number of weight functions and intervals the set of orthogonal polynomials is known. Stroud and Secrest [13] give x_i 's and A_i 's for a variety of weight functions and internals. The degree of the formula determines the number of x_i 's and A_i 's. A 2M-1 degree formula will have M points and is exact for polynomial integrands of degree 2M-1 or less.

Pierce applied Gaussian quadrature integration to two dimensions where the solution is of the form

$$ffg(x,y)dxdy = \sum_{i,j} \sum_{j} D_{ij}g(x_{ij},y_{ij})$$

The integration is over an annulus in the x,y plane with inner radius R and outer radius 1. Rewriting the integral in polar coordinates

$$\int_{R}^{1} \int_{0}^{2\pi} rG(r,\theta) d\theta dr \approx \sum_{i,j} \sum_{i,j} G(r_{j},\theta_{i})$$
 (5.14)

Pierce showed the summation could be rewritten as

where A_i and B_j are the weights for the radial and angular directions respectively and 4M+3 is the degree of accuracy.

For an arbitrary annulus the formula may be obtained by rewriting equation (5.14)

$$I = \int_{0}^{2\pi} \int_{r_1}^{r_2} f(r,\theta) r dr d\theta$$
 (5.16)

and letting

$$\rho = \frac{r}{r_2}$$

$$I = r_2^2 \int_0^{2\pi} \int_{r_1/r_2}^{1} \rho f(r_2 \rho, \theta) d\rho d\theta.$$
 (5.17)

The approximate solution can be derived as

$$I = \sum_{i=1}^{4 (m+1)} \sum_{j=1}^{m+1} A_{i}B_{j}f \left(\sqrt{r_{1}^{2} + \sigma_{j}(r_{2}^{2} - r_{1}^{2})}, e_{i} \right)$$
 (5.18)

where

$$A_i = 2\pi/4 (m+1)$$

and

$$B_j = w_j (r_2^2 - r_1^2)/2, w_j = weight$$

M+1 is the order of the orthogonal polynomial, in this case the Legendre polynomial on the interval (0,1) and σ_j 's are the zeros of this polynomial. This type of formula is known as a spherical product Gauss formula [12].

The annuli were picked at radii corresponding to the zeros of $J_{o}(z)$ because $J_{o}(z)$ is the fastest damping component

of the integrands. One additional annulus inside the first zero was found necessary to account for the $\frac{1}{\alpha+\chi^2}$ term when α is small and χ (and χ) approach zero. The distances between annuli outside the first five to eight zeros were found not to be critical and thus spaced arbitrarily.

The Gaussian Quadrature formula used was a Gauss-Legendre 24-point formula as listed in [13] thus giving a degree of accuracy of 99. The integration is limited to only the first quadrant as mentioned below. Thus for a 24-point formula, (M+1)² or 576 points per annulus were used to evaluate the integral.

The expression for β_n is easily seen to be an even function both radially and about both planer axes, therefore, only integration over the first quadrant was necessary. With a little more difficulty, $F_n(\hat{r}_0)$ and F_n can be shown (see Appendix E) to exhibit the same property provided one expression is used for even n and another for odd n. By calculating only over the first quadrant computer time was considerably shortened.

In order to solve for coefficients of arbitrary order, the Bessel functions of that order must first be obtained. The available library subroutines were found inadequate for this use for two reasons. First, these routines have an upper limit of 100 on the order of the Bessel function that can be computed and second, the routine had to be called for each individual order.

The recursive relation

$$J_n(y) = \frac{2n}{y} J_{n+1}(y) - J_{n+2}(y)$$
 (5.19)

provides a basis for generating an array of Bossel functions of various orders. However for accuracy $J_{n+1}(y)$ and $J_{n+2}(y)$ must be known.

The most accurate method to numerically calculate Besse! functions of various orders and arguments was found to be a uniform asymptotic expansion involving Airy functions [1]:

$$J_{n}(nz) = \left(\frac{4 \Lambda}{1-z}\right)^{\frac{1}{4}} \left\{ \frac{A_{i}(n^{2/3}\Lambda)}{n^{1/3}} \sum_{k=0}^{\infty} \frac{a_{k}(\Lambda)}{n^{2k}} + \frac{A_{i}!(n^{2/3}\Lambda)}{n^{5/3}} \sum_{k=0}^{\infty} \frac{b_{k}(\Lambda)}{n^{2k}} \right\}$$
(5.20)

where \mathbf{A}_{i} and \mathbf{A}_{i}^{*} are two types of Airy functions and

The Airy functions are calculated using the equations

$$A_{i}(x) = \frac{1}{2} x^{-i} e^{-\delta} f(-\delta)$$

and

$$A_{i}^{t}(x) = -\frac{1}{2} x^{i_{i}} e^{-\delta} g(-\delta)$$

where

$$\delta = \frac{2}{3} x^{3/2} \tag{5.22}$$

and

$$x = n^{2/3} \Delta$$

The functions $f(-\delta)$ and $g(-\delta)$ are tabulated and linear interpolation is used to determine their values. A is guaranteed to always be positive if n is selected to be larger than the argument.

By picking an order much higher than the argument, ν , and $J_{\nu-1}$, using the asymptotic expansion, the recursive relationship (5.19) may be used to generate an array of successive orders of the Bessel function down to and including J_0 . However, error builds up rapidly using this method. A second relation

$$1 = J_0(y) + 2J_2(y) + 2J_4(y) + \dots$$
 (5.23)

may be used to generate a normalizing factor

$$C = 1/[J_0(y) + 2J_2(y) + 2J_4(y) + ...]$$
 (5.24)

Multiplying each Bessel function value generated in the recursive relation by C resulted in very accurate values being obtained. The computer program used to solve the equations is listed as Appendix H.

To evaluate the numerical results a check was used which summed the coefficients and comprised this summation to an analytic expression for a sum that could easily be solved numerically.

For a particular phase angle of the nutation cycle, $\omega_0 t$, the radiance function for a target is

$$H(\omega_{o}t) = \sum_{k}^{\infty} H_{k}(\vec{r}_{o})e^{jk\omega_{o}t}$$
 (5.25)

For a point target and a rectangular, circularly nutating detector

$$H(\omega_{o}t) = \int e^{-2\pi^{2}\sigma^{2}k_{x}^{2}} \frac{\sin\pi_{w}k_{x}}{\pi k_{x}} e^{j2\pi k_{x}(\rho\cos\omega_{o}t - (r_{x}-x_{o}))} dk_{x}$$

$$\cdot \int e^{-2\pi^{2}\sigma^{2}k_{y}^{2}} \frac{\sin\pi hk_{y}}{\pi k_{y}} e^{j2\pi k_{y}(\rho\sin\omega_{o}t - (r_{y}-y_{o}))} dk_{y}$$
(5.26)

which has the solution (see Appendix F)

$$H(\omega_{o}t) = \frac{1}{4} \left[F_{1}(r_{x}, x_{o}, w) - F_{1}(r_{x}, x_{o}, w) \right] \left[E_{2}(r_{y}, y_{o}h) - F_{2}(r_{y}, y_{o}, h) \right]$$

where

$$E_{1}(r,z,\alpha) = Erf\left[\frac{r-z-\rho\cos\omega_{o}t}{\sqrt{2}o} + \frac{\alpha}{2\sqrt{2}\sigma}\right]$$

$$E_{2}(r,z,\alpha) = Frf\left[\frac{r-z-\rho\sin\omega_{o}t}{\sqrt{2}\sigma} + \frac{\alpha}{2\sqrt{2}\sigma}\right]$$

and Erf is the error function

$$Erf(x) = \frac{2}{\sqrt{\pi}} \int_{0}^{x} e^{-u^2} du$$
 (5.27)

By evaluating the analytic expression at the phase instant (ω_0 t) that the detector crosses the target and comparing

it against equation (5.27) for an appreciably high n, a reasonable check was made on the accuracy of the individual signal coefficients.

For the coefficients averaged over \vec{r}_o , the check was made at $\omega_o t$ = 0 where

$$H(o) = \sum_{k}^{\infty} \tilde{H}_{k}$$
 (5.28)

and the analytic expression is

$$H(o) = \frac{1}{4} \left[E_1(r_x, k_o, w) - E_1(r_x, x_o, -w) \right] \left[E_3(r_y, y_o h) - E_3(r_y, y_o, h) \right]$$

where

$$E_3(r,z,\alpha) = F_2(r,z,\alpha)|_{\omega_0 t=0}$$
 (5.29)

The rms background voltage was used to check the coefficients in the correlation function. The mean square voltage is

$$\bar{\mathbf{e}}_{B}^{2} = E\{B^{2}(t)\} = R_{B}(0) = \sum_{k=0}^{\infty} B_{k}$$
 (5.30)

Appendix B shows that the correlation function may be written as

$$R_{B}(\tau) = \int |\tau(k)|^{2} W_{B}(k) J_{O}\left(4\pi\rho k \sin \frac{\omega_{O}\tau}{2}\right) d^{2}k \qquad (16B)$$

Substituting (5.1), (5.2), and (5.36) in (16B) and evaluating at τ = 0

$$\frac{\bar{e}_{B}^{2} = \sigma_{B}^{2\alpha} \int \frac{\sin^{2}(\pi k_{x}^{W})}{(\pi k_{x})^{2}} \frac{e^{-4\pi^{2}\sigma^{2}k_{x}^{2}}}{\sigma^{2} + (2\pi k_{x})^{2}} dk_{x}$$

$$\frac{\sigma_{B}^{2\beta} \int \frac{\sin^{2}(\pi k_{y}^{W})}{(\pi k_{y})^{2}} \frac{e^{-4\pi^{2}\sigma^{2}k_{y}^{2}}}{\sigma^{2} + (2\pi k_{y})^{2}} dk_{y} . \quad (5.31)$$

It can be shown (see Appendix F) that an analytic expression is

$$\bar{e}_B^2 = 4\alpha\beta\sigma_B^2R(w;\alpha)R(h;\beta)$$

where

$$R(\mu,\alpha) = \frac{-\mu}{\alpha^2} \left\{ 1 - \operatorname{Erfc} \left(\frac{2\mu}{\sigma} \right) - \frac{2\sigma}{\sqrt{\pi}\mu} \left[1 - e^{-\left(\frac{\mu}{2\sigma} \right)^2} \right] \right\}$$

$$+ \frac{e^{\alpha^2 \sigma^2}}{2\alpha^3} \left[e^{\alpha\mu} \operatorname{Erfc} \left(\alpha\sigma + \frac{\mu}{2\sigma} \right) \right]$$

$$+ e^{-\alpha\mu} \operatorname{Erfc} \left(\alpha\sigma - \frac{\mu}{2\sigma} \right) - 2\operatorname{Frfc} (\alpha\sigma)$$
(5.5.)

and Erfc is the complimentary error function

$$\operatorname{Eric}(x) = \frac{2}{\sqrt{\pi}} \int_{-\infty}^{\infty} e^{-u^2} du . \qquad (5.33)$$

The computer program used to solve equations (5.6), (5.7) and (5.9) and calculate the infinite summations is listed as Appendix H.

VI. NUMERICAL RESULTS

Numerical results constitute solving for the signal and background coefficients and using these coefficients to specify the form of the filter to be used in the threshold detector as well as to calculate the probability of detection as a function of signal-to-noise ratio (SNR), background-to-noise ratio (BNR) and number of nutations (M) used to make the decision. The probability of detection is determined for the averaged equivalent signal-to-noise ratio and for specific point targets.

Many different values for system parameters and target locations were used to compute the coefficients. A typical set of parameters is 1 :

nutation radius = 15
width of detector = 30
height of detector = 1
blur circle standard deviation = .5
background correlation length x-direction = 20
background correlation length y-direction = 20
position of detector in detector coordinate system
 x direction = 15
 y direction = 0
target coordinates (if desired)
 x direction = 0
 y direction = 0
pointing error standard deviation = 7.5

A detector utilizing the parameters above would trace out the area in the image plane shown in Figure 6.1.

¹ All units are in milliradians.

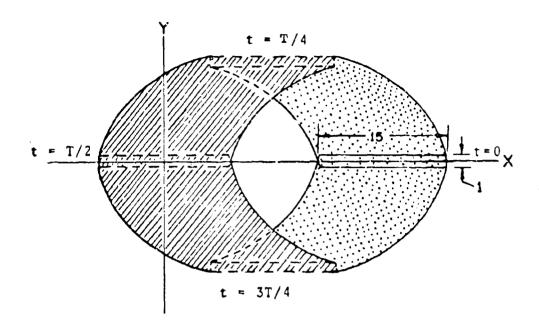


Figure 6.1. Scanned Area of Nutating Detector.

This type scan was chosen because it crossed the center of the coordinate system which would be the center of the target blur circle for zero pointing error. Also, in the case of a random pointing error in the absence of bias zero mean is the center of the coordinates. This is important because the threshold detector design is based on a Gaussian pointing error with zero mean.

A sample calculation of 121 coefficients using the parameters above is listed in Appendix G. The equations shown in Part V were used to check the accuracy of the coefficients. Table 6.1 lists typical values for the summation of 121 coefficients for various sets of parameters and the calculated summation values.

TABLE 6.1

	Calculated Sum	Actual Sum
Signal (point target)	.6826894	.6826993
	.6826895	.6826994
	.5438002	.5438087
Signal (averaged)	.2990664D-6	.2991232D-6
	.2328138D-4	.2328142D-4
	.2651595D-1	.2651595D-1
Background	.6013477	.6012230
	.2347420	.2347420
	.8335525	.8319764

Besides the general shapes of the background and signal spectra, variations in the spectra due to detector size, background correlation lengths and coordinates of the detector are of interest. Curves in the following figures show the effects of these variations.

Figures 6.2 through 6.5 show the envelope of the one side of the two-sided background power spectral density

$$S_{B}(\omega) = \sum_{k}^{\infty} \beta_{k} \delta(\omega - k\omega_{0})$$

where ω_0 is normalized to 2π . The curves of Figures 6.2 and 6.3 illustrate the effect of varying the background correlation lengths. In Figure 6.2 the correlation lengths are the same in both x- and y-directions. Practical measurements, however, indicate that correlation lengths in the horizontal (x-) direction may be longer than in the vertical (y-) direction. The curves in Figure 6.3 are for this case.

The curves of Figure 6.4 and 6.5 relate the spectrum to changes in size of the detector. In Figure 6.4, only the detector with BNR equal to 9 and number of nutations equal to 100 are shown in Figures 6.11 through 6.13. Each figure shows the envelope of one side of a double sided spectrum.

Two cases are considered. First, threshold detection systems designed for different Gaussian pointing errors but with the same size detector are shown in Figures 6.11 and 6.12 and second, detection systems designed for the same printing error but utilizing different size detectors are shown in Figures 6.12 and 6.13.

The curves shown in Figures 6.14 through 6.16 show the matched filter to a particular point target located within the scan. Notice each of the filters exhibits a band pass characteristic. This is because of the additional noise component (background). The oscillations observed in the point target spectra of Figures 6.6 to 6.8 have been suppressed to provide a look at realizable filters.

The probability of detection for the threshold detector was shown in Part IV to be

$$Q_d = SNR \sqrt{M} \left[\sum_{k=1}^{\infty} |\bar{f}_k|^2 \frac{1}{1 + MQ_k BNR^2} \right]^{\frac{1}{2}}$$

Figures 6.17 through 6.20 show Q_{d} as a function of SNR, BNR and M.

The probabilities of detection for point targets located at specific points in the scan have also been plotted. The probability of detection for a specific point target is shown as the

width of the detector is varied. As width is increased, so is the area scanned by the detector. Figure 6.5 shows the difference in spectra between a rectangular detector and a square detector which have approximately the same surface area. It should be obvious however that the rectangular detector scans much more area per nutation than a square detector of identical surface area.

The envelopes of signal coefficients (magnitude only) for a point target are plotted in Figures 6.6 through 6.8. Phase information depends only on the location of the target relative to the initial point for the nutation cycle which is along the x axis. This has not been plotted. Changes in the magnitude due to changes in detector size are shown.

Figures 6.9 and 6.10 show the envelopes of coefficients for the signal that is averaged with respect to a Gaussian pointing error. The bandwidth of the averaged signal is much less than a point target as expected. The figures show the spectra for two different standard deviations of pointing error.

Matched filters (magnitude only) for the threshold

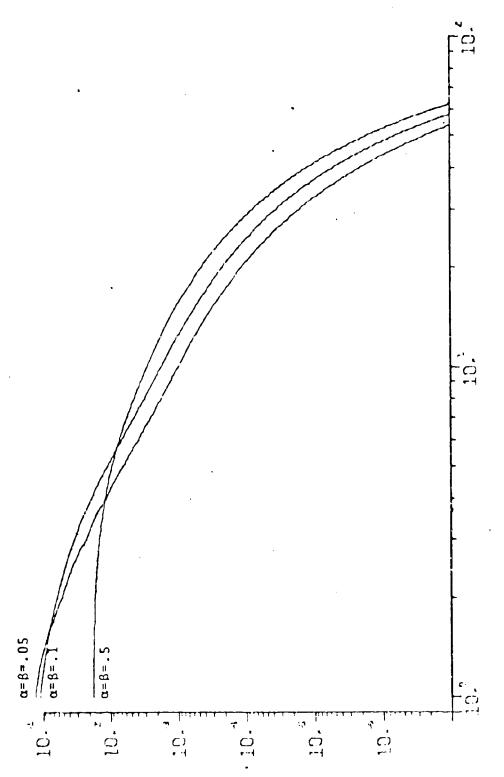
$$Q_{d}^{*} = SNR^{*}\sqrt{M} \frac{\sum_{k} f_{k}(\vec{r}_{o}) f_{k}^{*} \frac{1}{1 + MQ_{k}BNR^{2}}}{\left[\sum_{k} |\vec{f}_{k}|^{2} \frac{1}{1 + MQ_{k}BNR^{2}}\right]^{\frac{1}{2}}}$$

where

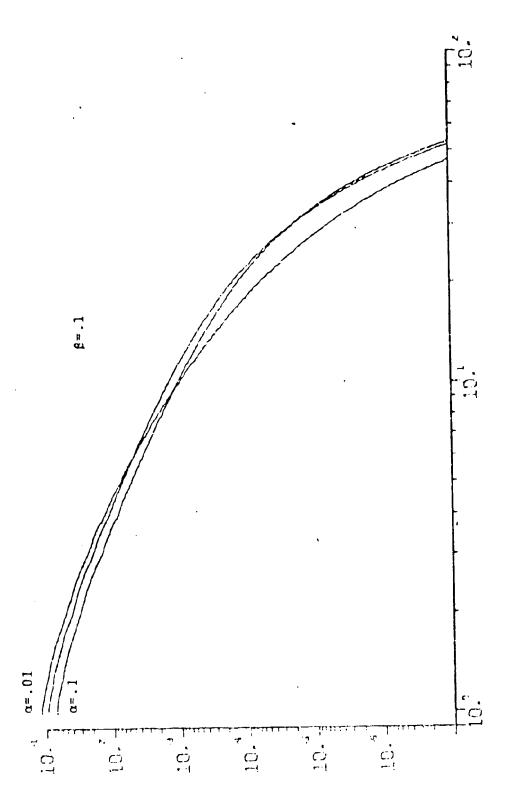
SNR' =
$$\frac{1}{M} \sum_{n=1}^{M} a_n \sqrt{T} / \sqrt{\frac{S}{2}}$$

Figures 6/21 through 6/24 show $Q_{\mbox{\scriptsize d}}^{\,\prime}$ as a function of SNR', BNR and M.

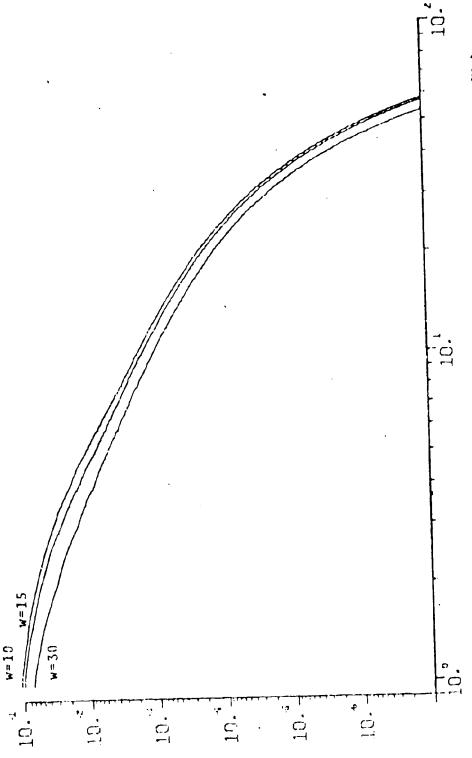
The parameters of the nutating optical system have been varied to show their effect on \textbf{Q}_d and $\textbf{Q}_d^*.$



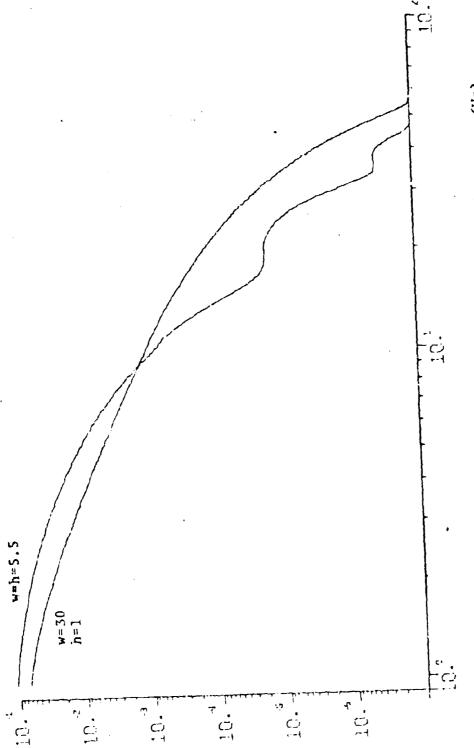
Background Power Spectral Density vs. Frequency (Hz). Detector parameters: RHO=10, SIG=.5, w=10, h=1 Figure 6.2.



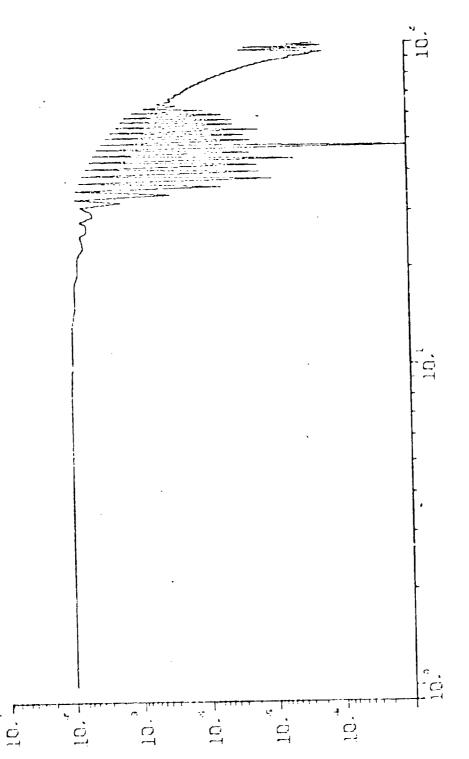
Background Power Spectral Density vs. Frequency (Hz). Detector parameters: RHO=15, SIG=.5, w=30, h=1 Figure 6.3.



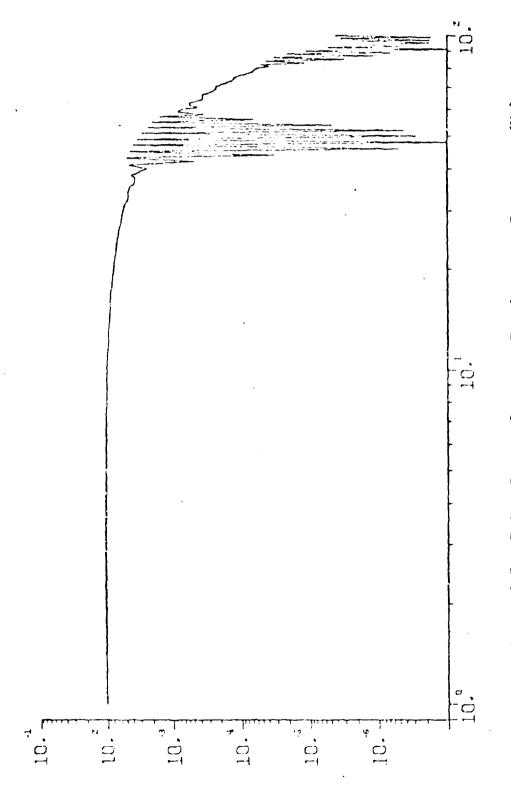
Background Power Spectral Density vs. Frequency (Hz). Common detector parameters: RH0=15, SIG=.5, h=1 Barhanund narameters: $\alpha=\beta=.1$ Figure 6.4.



Background Power Spectral Density vs. Frequency (Hz) Common detector parameters: RHO=15, SIG=.5 Background parameters: $\alpha=\beta=.1$ Figure 6.5.

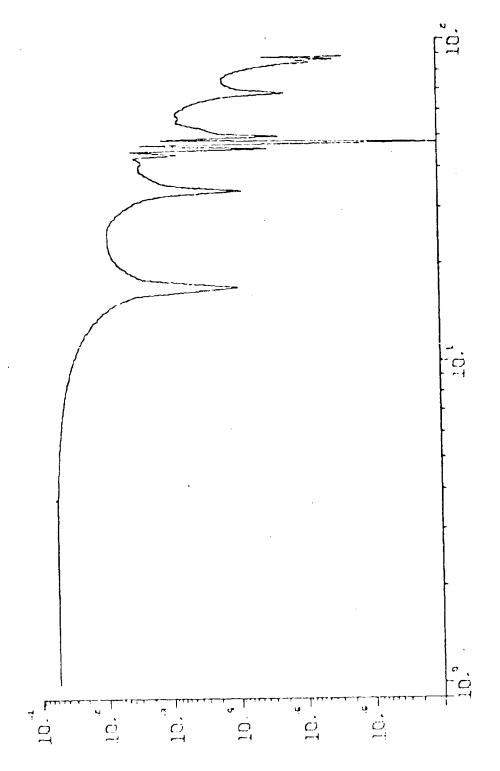


Point Target Spectrum Envelope vs Frequency (Hz) Detector parameters: RHO=15, SIG=.5, w=30, RHO=15, SIG=.5, w=30, h=1, $x_0=15$, $y_0=0$ Target location: Figure 6.6.



Point Target Spectrum Envelope vs. Frequency (Hz). Def.... narameters: RHO=15, SIG=.5, w=10, h=1, $x_0=15$, $y_0=0$ Figure 6.7.

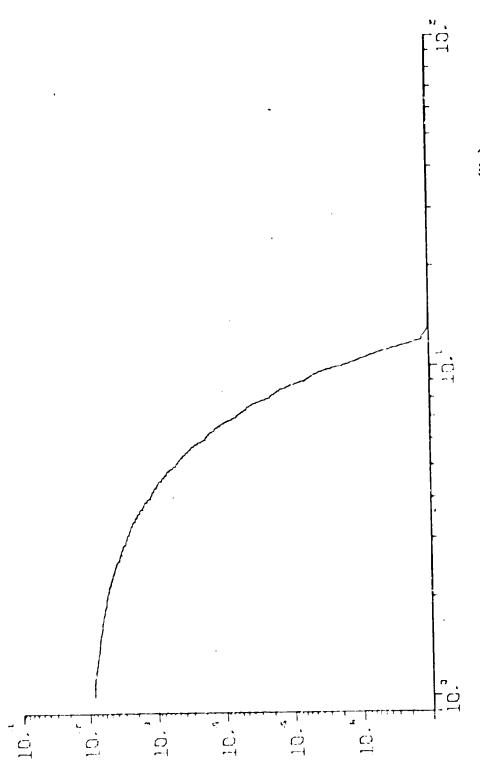
Target location: $R_x = 0$, $R_y = 0$



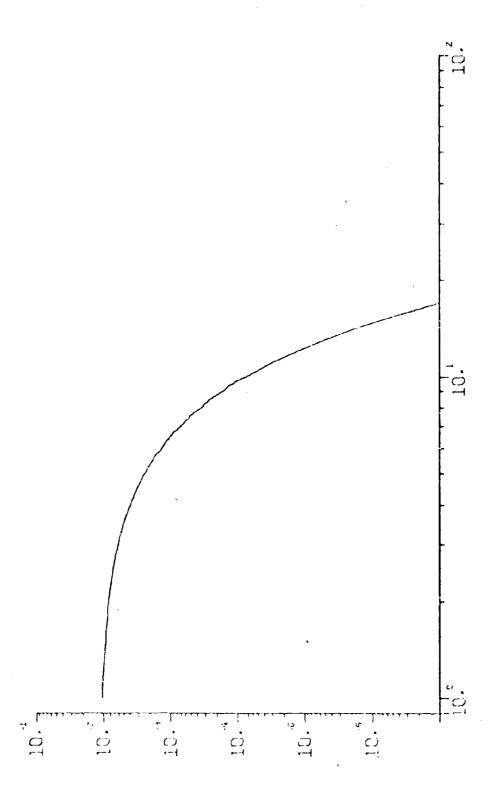
Point Target Spectrum Envelope vs. Frequency (Hz) Detector parameters: Figure 6.8.

Target parameters:

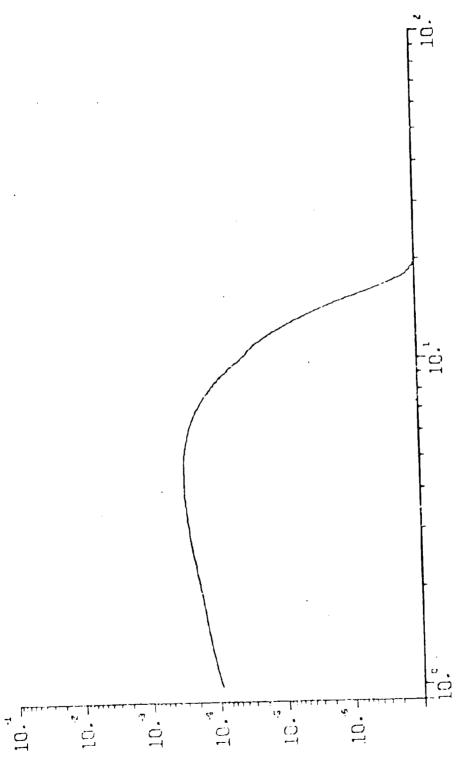
68



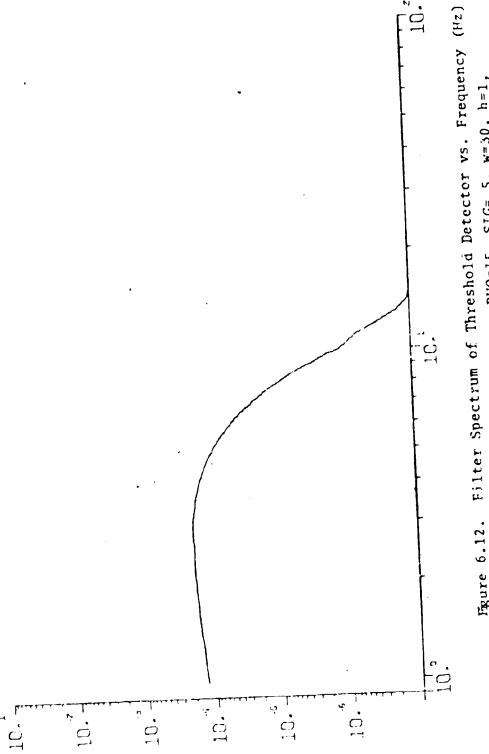
Drtector parameters: RHO=15, SIG=5, w=30, h=1, $x_0=15$, $y_0=0$ Std. Dev. of Gaussian Pointing error = 7.5 Averaged Signal Spectrum vs. Frequency (Hz). Figure 6.9.



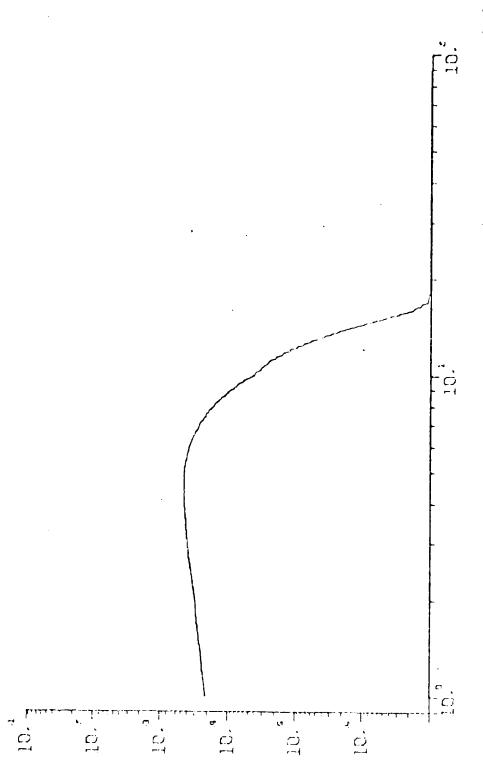
Averaged Signal Specirum Envelope vs. Frequency. Fig=15, SIG=.5, w=30, $\hat{n}=1$, $x_0=15, y_0=1$ Std. Dev. of Gaussian Pointing error = Detector parameters: Figure 6.10.



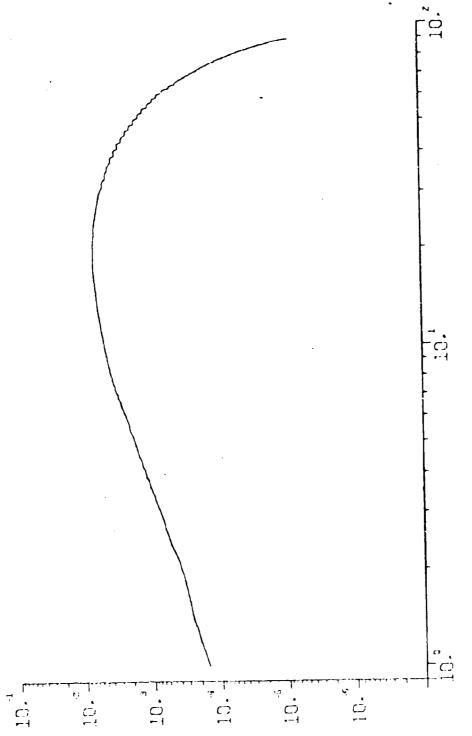
Filter Spectrum for Threshold Detector vs. Frequency (Hz) RHO=15, SIG=.5, w=30, h=1, $x_0=15$, $y_0=0$ Std. Dev. of Gaussian Pointing error = 5.0 Detector parameters: Figure 6.11.



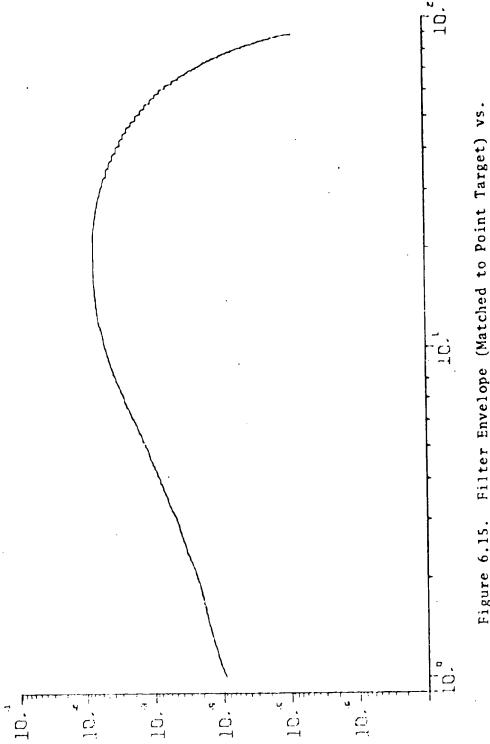
Detector parameters: RHO=15, SIG=.5, w=30, h=1, $x_0=15$, $y_0=0$ Std. Dev. of Gaussian pointing error = 7.5 Figure 6.12.



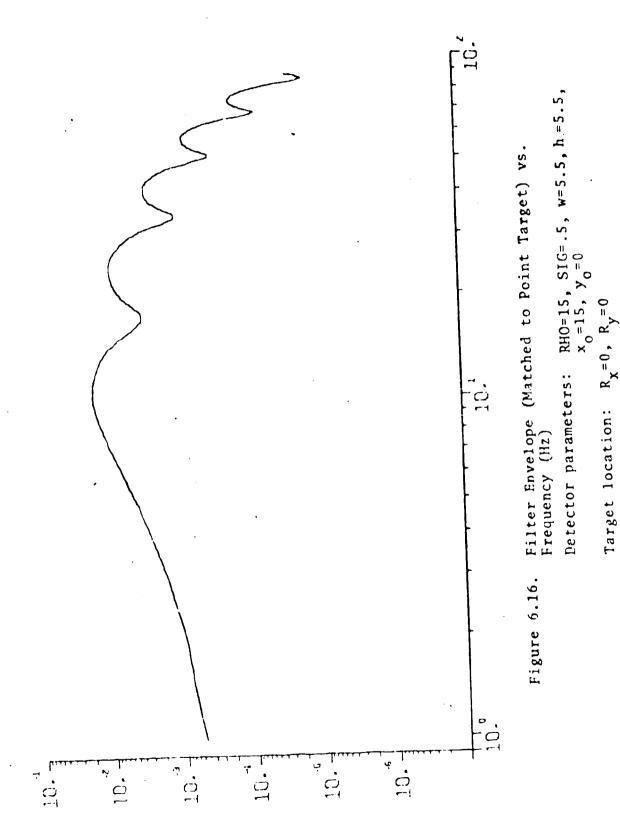
Filter Spectrum for Threshold Detector vs. Frequency (Hz) Detector parameters: RHO=15, SIG=.5, w=5.5, h=5.5, $x_9=15$, $y_0=0$ Std. Dev. of Gaussian pointing error = 5.0 Figure 6.13.

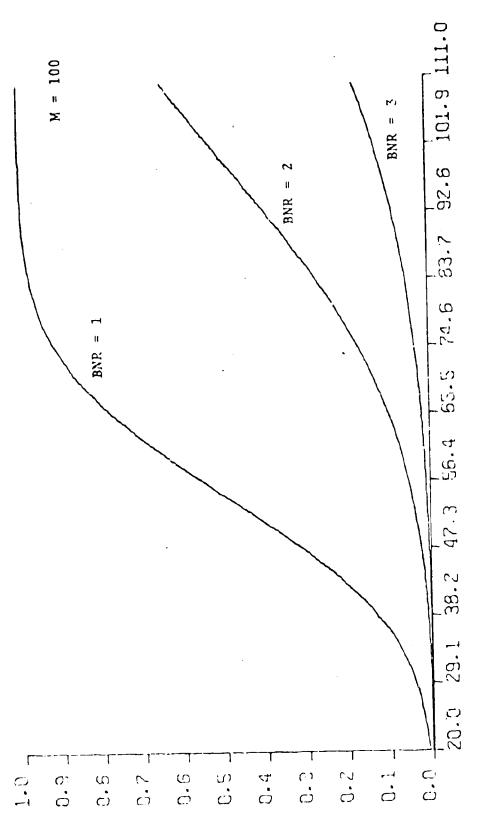


Target location: $R_{\rm X}=15$, SIG=.5, w=30, h=1, $N_{\rm O}=15$, $N_{\rm O}=0$ Filter Envelope (Matched to Point Target) vs. Frequency (Hz). Figure 6.14.

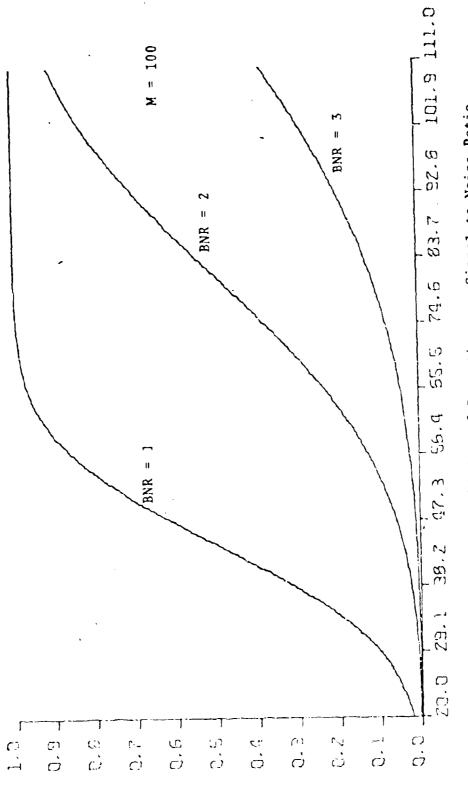


Detactor narameters: RHO=15, SIG=.5, w=10, h=1, x_0=15, y_0=0
Target location: $R_x=0$, $R_y=0$ Filter Envelope (Matched to Point Target) vs. Frequency (Hz) Figure 6.15.



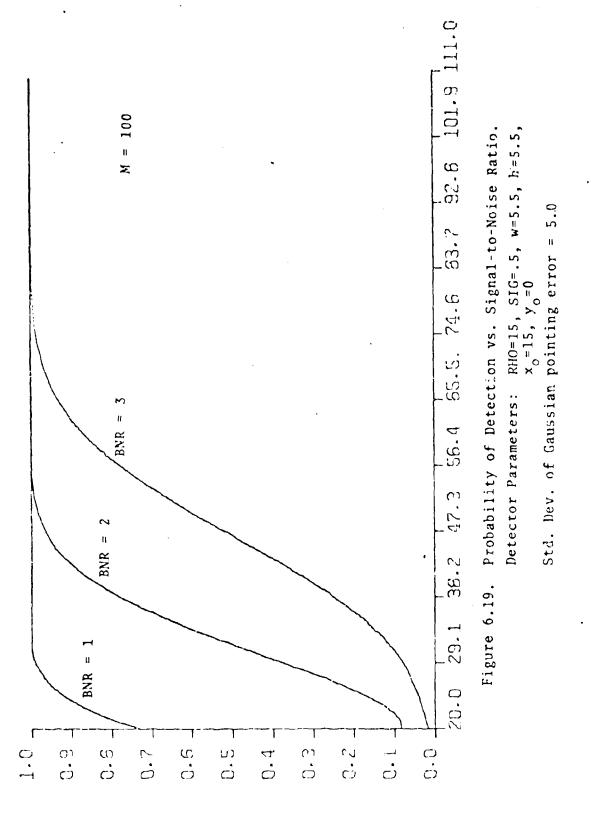


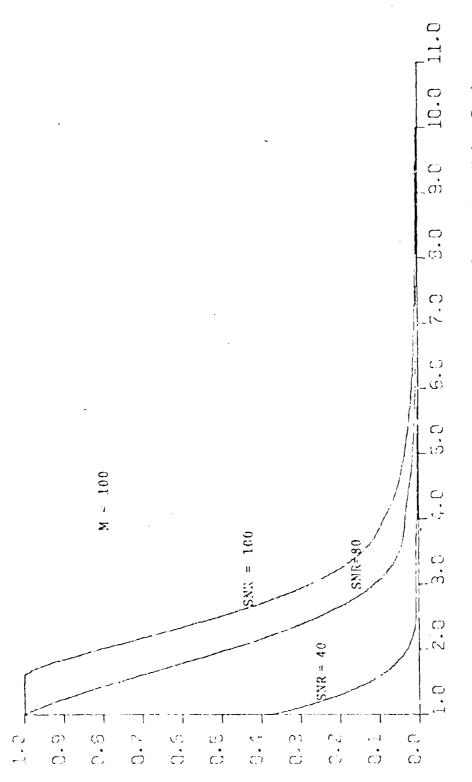
Probability of Detection vs. Signal-to-Noise Ratio. RH0=15, SIG=.5, w=10, h=1, $x_0=15$, $y_0=0$ Std. Dev. of Gaussian pointing error = 5.0Detector parameters: Figure 6.17.



Probability of Detection vs. Signal-to-Noise Ratio. RHO=15, SIG=.5, w=30, h=1, $x_0=15$, $y_0=0$ Detector parameters: Figure 6.18.

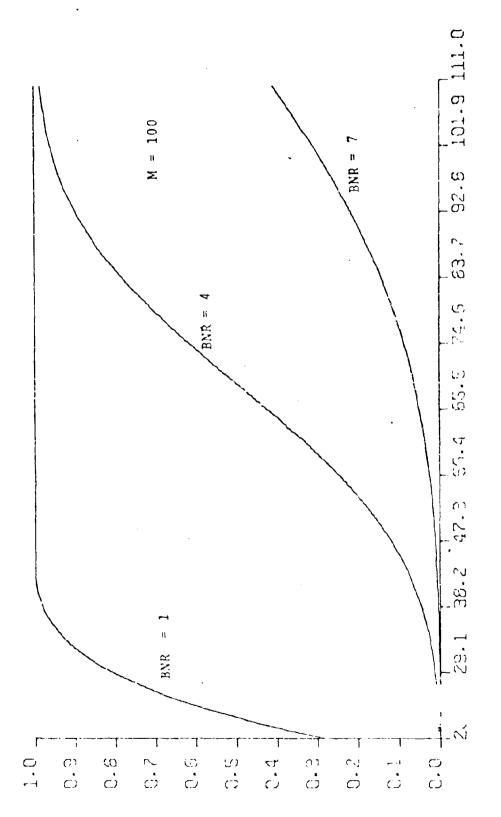
Std. Dev. of Gaussian pointing error = 7.5



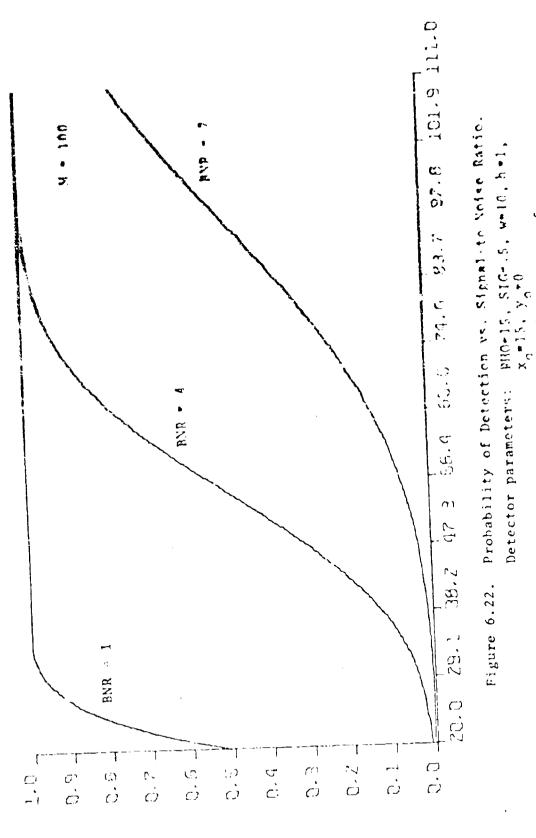


Probability of Detection vs. Background-to-Noise Ratio. Figure 6.20.

Detector parameters: RHO=15, SiG=.5, w=50, h=1, $x_0=15$, $y_0=0$ Std. Dev. of Gaussian pointing error = 7.5



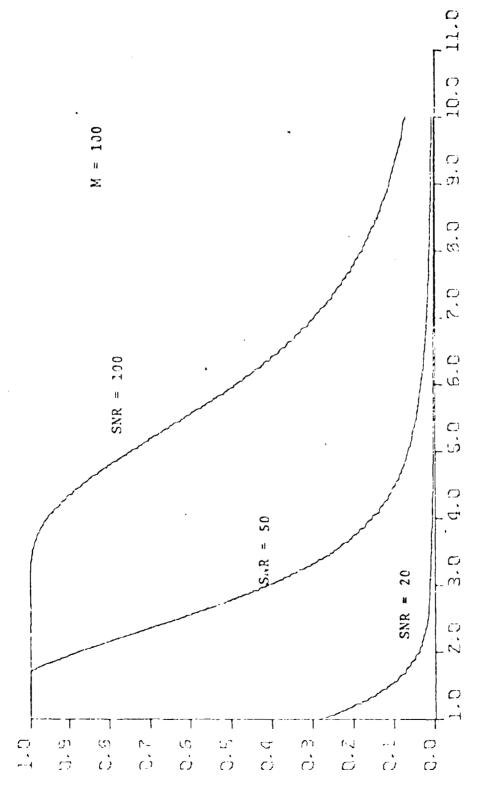
Probability of Detection vs. Signal-to-Noise Ratio. Detector Parameters: RHO=15, SIG=.5, w=30, h=1, x_0 =15, y_0 =0 Std. Dov. of Gaussian pointing error # 7 Target location: Ry+0, Ry+0 Figure 6.21.



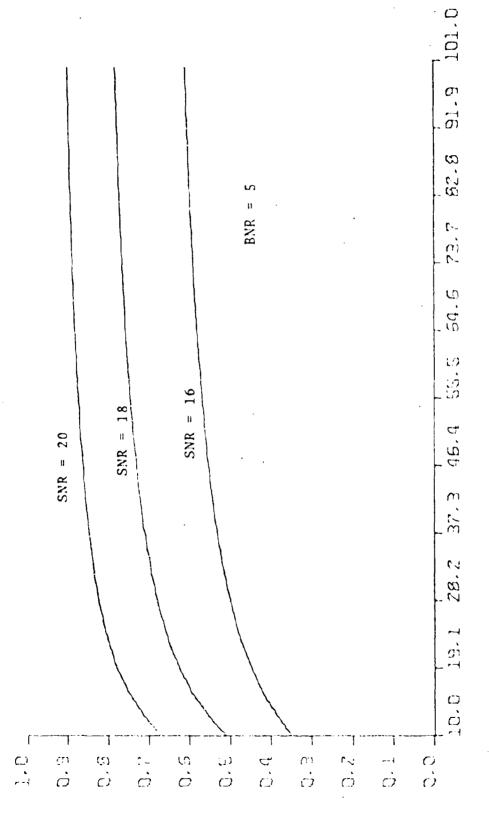
Std. Dev. of Gaussian pointing error " Target location: Rymo, Rymo

Detector parameters:

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RHO=15, SIG=.5, w=30, h=1, x_0 =15, y_0 =0 7.5 Probability of detection vs. Background-Std. Dev. of Gaussian pointing error = Target location = $R_X = 0$, $R_Y = 0$ Detector parameters: to-Noise Ratio. Figure 6/23.



Detector parameters: RHO=15, SIG=.5, w=5.5, h=5.5, $x_0=15$, $y_0=0$ Probability of Detection vs. Number of Nutations. Std. Dev. of Gaussian pointing error = 5 Target location = $R_x = 0$, $R_y = 0$ Figure 6.24.

VII. CONCLUSIONS

The problem investigated in this thesis was to determine an optimum processor to be used in detecting targets with an infrared nutating system. The nature of the system divided the problem into two areas: optimizing a spatial processor according to size and shape and optimizing a temporal processor that takes the output time varying voltage and decides if a target is actually present.

General equations for the output from a nutating detector were known from Samuelsson's work, however these equations had not been solved for a specific detector. At the same time it was observed that the form of the covariance function for the noise offered an easy solution to integral equation of the Karhunen-Loeve expansion which made statistical detection theory appealing. Some work had been done on the form of temporal processor using statistical detection theory but this was limited to only one nutation because of the common background noise component between nutations.

Harger's derivations provided a means for describing a temporal processor using statistical detection theory that based its decisions on multiple observations and included the background noise. His work was extended here to include unknown parameters, amplitude and position, which more closely characterized the system. This extension, called a threshold detector, was shown to be optimum in the case of

small signal-to-noise ratios where detection is most difficult. The derivation also used the Gaussian assumption to describe the background. This may or may not be correct. But in practice, the detector's performance may not suffer greatly if this assumption is wrong.

To implement the threshold detector, the integrals describing the detecor output were solved using the Gaussian quadrature method of numerical integration. Checking the computed coefficients by a summation provided a means of accepting the validity of the integrations. The form of the threshold detection system using a rectangular detector was determined and the frequency spectrum of the optimum filter was shown. To specify performance, the probability of detection is plotted against signal-to-noise and background-to-noise ratios and the number of nutations on which the decision was based. The signal spectrum for a point target and background power spectral density is also plotted.

A designer may find these results useful in developing a system or investigating the performance of an actual system s compared to the optimum. He may also use these results to determine the size of a rectangular spatial filter to be used.

Future work should include extending these results to spatial filters that are circular or elliptic. The same numerical integration algorithm could probably be used. The threshold detector was found to be highly sensitive to background correlation lengths. Thus, some form of adaptive processor should be included to minimize the effect of

background. At the present time, a measurement program is underway at Naval Weapons Center, China Lake, California to determine average background correlation lengths but for different environments, these correlation lengths could be expected to vary considerably.

APPENDIX A

DERIVATION OF SIGNAL SPECTRAL COFFFICIENTS

The radiance on the detector is periodic because of the nutation therefore it can be expressed in a Fourier series,

$$H(t) = \sum_{n=0}^{\infty} H_{n} e^{jn\omega_{0}t}$$
(1A)

where

$$H_{n} = \frac{1}{T} \int_{0}^{T} H(t) e^{-jn\omega_{0}t} dt$$
 (2A)

But,

$$H(t) = \int N'(\underline{r}) \tau(\underline{r} \cdot \underline{\rho}(t)) d^2 \underline{r} . \qquad (3A)$$

Substituting (3A) in (2A)

$$H_{n} = \int d^{2} \mathbf{r} \, N'(\mathbf{r}) \, \frac{1}{T} \int_{0}^{T} d\mathbf{r} \, \tau(\mathbf{r} \cdot \mathbf{r}) (\mathbf{t}) \, e^{-j n \omega_{0} \mathbf{t}}$$
(4A)

The transmittance, τ , may be expressed in terms of the inverse Fourier transform, assuming an infinite image plane,

$$\tau(\mathbf{r} \cdot \varrho(\mathbf{t})) = \int d^2k \ \tau^*(k) e^{-j2\pi k \cdot (\mathbf{r} \cdot \varrho(\mathbf{t}))}. \tag{5A}$$

Rewriting the exponent,

$$k \cdot (r - \rho(t)) = k \cdot r - \rho k_x \cos \theta t - \rho k_y \sin \theta t$$
 (6A)

Equation (4Λ) can be rewritten,

$$H_{n} = d^{2} r N'(r) \frac{1}{T} \int_{0}^{T} dt \int d^{2}k \tau^{*}(k) e^{-j2 k \cdot r}$$

$$= \int_{0}^{1} 2\pi A \sin(\omega_{0} t + \theta) e^{-jn\omega_{0} t}$$

where

$$A = \sqrt{k_x^2 + k_y^2}$$

$$\theta = \tan^{-1} k_x/k_y.$$
(7A)

Observe that

$$N'(k) = \int d^2 r N'(r) e^{-j 2\pi k \cdot r}$$
 (8A)

thus

$$H_{n} = \int d^{2}k \, N'(k) \tau^{*}(k) e^{jn\theta} \, \frac{1}{T} \int_{0}^{T} dt \, e^{-jn(\omega_{0}t+\theta)} \cdot e^{jnAsin(\omega_{0}t+\theta)}$$
(9A)

Using the relation

$$J_{n}(z) = \frac{1}{T} \int_{0}^{T} e^{-jn\phi} e^{jz\sin\phi} d\phi, \qquad (10A)$$

the expression for the coefficients becomes

$$H_n = \int d^2k \, N'(k) \, \tau^*(k) \, e^{j\,n\theta} \, J_n \left(2\pi\rho \sqrt{k_x^2 + k_y^2}\right).$$
 (11A)

Let

$$\varphi = \tan^{-1} ky/kx$$

then

$$j^n e^{-jn\phi} = e^{jn\theta}$$

Rewriting (11A)

$$H_{n} = j^{n} \int d^{2}k \, N'(k) i^{*}(k) e^{-jn\phi} J_{n} \left(2\pi\rho \sqrt{k_{x}^{2} + k_{y}^{2}}\right). \quad (12A)$$

APPENDIX B

DERIVATION OF BACKGROUND CORRELATION COEFFICIENTS

The $n^{\mbox{th}}$ power spectrum component of the background at frequency $n\omega_{\mbox{\scriptsize 0}}$ is given by

$$\beta_{n} = \frac{1}{T} \int_{0}^{T} R_{B}(\tau) e^{-jn\omega_{0}} d\tau$$
 (1B)

But $R_{\underline{p}}(\tau)$ is the correlation function of detector output averaged with respect to the starting time.

$$R_{B}(\tau) = \frac{1}{T} \int_{0}^{T} E[B(t+\tau)B(t)] dt$$
 (2B)

Using the expression for B(t) from Part II,

$$R_{B}(\tau) = \frac{1}{T} \int_{0}^{T} dt \ E\left[\int N'(\underline{r}) \tau(\underline{r}_{1} - \varrho(t + \tau)) d^{2}\underline{r}_{1}\right]$$

$$\cdot \int N'(\underline{r}_{2}) \tau(\underline{r}_{2} - \varrho(t)) d^{2}\underline{r}_{2} \qquad (3B)$$

Interchanging the order of expectation and integration,

$$R_{\mathbf{B}}(\tau) = \frac{1}{T} \int_{0}^{T} d\tau \int d^{2} \mathbf{r}_{1} \int d^{2} \mathbf{r}_{2} \tau \left[\mathbf{r}_{1} - \varrho(t + \tau) \right]$$

$$+ \tau \left[\mathbf{r}_{2} - \varrho(t) \right] E\left[N'(\mathbf{r}_{1}) N'(\mathbf{r}_{2}) \right]$$
(4B)

Assuming stationary background,

$$E[N'(\underline{r}_1)N'(\underline{r}_2)] = \sharp_B'(\underline{r}_1 - \underline{r}_2)$$
 (5B)

where $\phi_B^i(r_1-r_2)$ is the background covariance function on the image plane.

It is now convenient to express the integrand in terms of two-dimensional spatial Fourier transform. To begin with

$$\tau(r) = \int \tau(k)e^{j2\pi k \cdot r} d^2k$$
 (6B)

where $\tau(\underline{k})$ is the Fourier transform of the aperture function $\tau(\underline{r}) \text{ and } \underline{k} = (k_x, k_y).$ Then

$$\tau \begin{bmatrix} \underline{r}_{1} - \varrho(t+\tau) \end{bmatrix} \tau \begin{bmatrix} \underline{r}_{2} - \varrho(t) \end{bmatrix}$$

$$= \int d^{2}\underline{k}_{1} \tau (\underline{k}_{1}) \int d^{2}\underline{k}_{2} \tau^{*} (\underline{k}_{2}) e^{j2\pi (\underline{k}_{1} \cdot \underline{r}_{1} - \underline{k}_{2} \cdot \underline{r}_{2})}$$

$$\cdot e^{-j2\pi \begin{bmatrix} \underline{k}_{1} & \varrho(t+\tau) - \underline{k}_{2} \varrho(\tau) \end{bmatrix}}$$
(7B)

Substituting (5B) and (7B) in (4B) and interchanging the order of integration,

$$R_{B}(\tau) = \int d^{2}k_{1} \ \tau(k_{1}) \int d^{2}k_{2} \ \tau^{*}(k_{2})$$

$$\cdot \frac{1}{1} \int_{0}^{T} dt \ e^{-j2\pi \left[k_{1} \cdot \varrho(t+\tau) \cdot k_{2} \cdot \varrho(\tau)\right]}$$

$$\cdot \int d^{2}r_{1} \ e^{j2\pi k_{1} \cdot r_{1}} \int d^{2}r_{2} \ \phi_{B}^{*}(r_{1} - r_{2}) e^{-j2\pi k_{2} \cdot r_{2}}$$
(8B)

The last integral becomes, with $R = r_1 \cdot r_2$,

$$\left[\int d^{2}\mathbf{R}^{\dagger}_{+}^{*}(\mathbf{R}) e^{-\mathbf{j} 2\pi \mathbf{k}_{2} \cdot \mathbf{R}}\right] e^{-\mathbf{j} 2\pi \mathbf{k}_{2} \cdot \mathbf{r}_{1}}$$

$$(98)$$

But the bracket above is the Fourier transform of $\frac{1}{B}(\mathbf{r})$, or the Wiener spectrum of background on the image plane. Let

$$W_{B}^{\dagger}(k_{2}) = \int : \frac{1}{B}(R) e^{-j2\tau k_{2} \cdot R} d^{2}R$$
(10B)

If $F_{O}(\underline{k})$ is the Fourier transform of the point spread function and $W_{B}(\underline{k})$ is the Wiener spectrum of background on the object plane, then

$$W_{\mathbf{B}}^{\dagger}(\underline{k}) = |F_{\mathbf{O}}(\underline{k})|^2 W_{\mathbf{I}}(\underline{k})$$
 (11B)

This relation is analogous to the output noise spectron expression due to a noise input to a linear system. In terms of $W_{\rm R}^{\star}(k)$ the last line of (8B) is

$$W_{B}'(k) \int_{C} \int_{C} \frac{j 2 \pi (k_{1} - k_{2}) \cdot r_{1}}{2} d^{2}r_{1} = W_{B}'(k) \delta(k_{1} - k_{2})$$
 (14B)

Substituting this into (8B) and performing the integration with respect to $\boldsymbol{k}_{2}\text{,}$

$$R_{B}(\tau) = \int d^{2}k |\tau(k)|^{2} W_{B}'(k)$$

$$\cdot \frac{1}{T} \int_{0}^{T} e^{j2\pi k \cdot [\rho(t+\tau) - \rho(t)]} dt \qquad (13B)$$

Since $\rho(t) = (\rho \cos_{-0} t, \rho \sin_{-0} t)$,

$$k \cdot [\varrho(t+\tau) \cdot \varrho(t)] = k_x e^{-cos\omega_0} (t+\tau) + k_y e^{-sin\omega_0} (t+\tau)$$

$$-k_x e^{-cos\omega_0} t - k_y e^{-sin\omega_0} t$$

$$= A(\tau) e^{-s\omega_0} t + B(\tau) sin\omega_0 t = e^{-s\omega_0} (e^{-s\omega_0} t - e^{-s\omega_0} t)$$

where

$$A(\tau) = \rho [k_x \cos \omega_0 \tau + k_y \sin \omega_0 \tau - k_x]$$

$$B(\tau) = \rho [-k_x \sin \omega_0 \tau + k_y \cos \omega_0 \tau - k_y]$$

$$C(\tau) = \sqrt{\Lambda^2 (\tau) + h^2 (\tau)} = \rho \sqrt{2(1 - \cos \omega_0 \tau)(k_x^2 + k_y^2)}$$

and

$$\psi(\tau) = \tan^{-1} \frac{B(\tau)}{\lambda(\tau)}$$
 (14B)

Therefore the last integral of (13B) is

$$\frac{1}{T} \int_{0}^{T} e^{j2\pi c \cos(\omega_0 \tau + \psi)} dt = J_0(2\pi c)$$
 (15B)

where $J_{\zeta}(\cdot)$ is the zeroth order Bessel function. Therefore, the correlation function $R_{B}(\tau)$ is given by

$$R_{B}(\tau) = \int |\tau(\underline{k})|^{2} W_{B}(\underline{k}) J_{o}\left(4\pi\rho\underline{k} \sin \frac{\omega_{o}\tau}{z}\right) d^{2}\underline{k}$$
 (16B)

where

$$\underline{k} = \sqrt{k_x^2 + k_y^2}$$
 and $d^2 \underline{k} = d\underline{k}_x d\underline{k}_y$.

Finally, the power spectrum coefficient β_n is, from (1B)

$$\beta_{n} = \int |\tau(\underline{k})|^{2} W_{B}^{*}(\underline{k}) d^{2}\underline{k}$$

$$\cdot \frac{1}{T} \int_{0}^{T} J_{o}\left(2a \sin \frac{\omega_{o}\tau}{z}\right) e^{-jn\omega_{o}\tau} d\tau \qquad (17B)$$

where a = $2\pi k\rho$. The last integral becomes with x = $\omega_0 \tau/2$,

$$\frac{1}{T} \int_{0}^{\pi} J_{0}(2a \sin x) \cos 2nx \ dx$$

$$- j \frac{1}{\pi} \int_{0}^{\pi} J_{0}(2a \sin x) \sin 2n\pi dx$$
 (18B)

The cosine integral equals $J_n^2(a)$ and the sine integral vanishes [2]. Therefore on using (18B)

$$\beta_{n} = \int |\tau(\underline{k})|^{2} |F_{o}(\underline{k})|^{2} |W_{B}(\underline{k})J_{n}^{2}(2\pi\rho\underline{k})d^{2}\underline{k}. \qquad (19B)$$

This is the general power spectrum expression for a nutating detector output. The power at frequency $n\omega_0$ is expressed in terms of aperture transform $\tau(\underline{k})$, optical transfer function $F_0(\underline{k})$, Wiener spectrum of background $W_B(\underline{k})$ and nutation radius through $J_n^2(2\pi\rho\underline{k})$. The integration is over the entire k-plane.

APPENDIX C

EXPECTATION OF Λ_{12}

The essential calculation required to find

$$\Lambda_{12}(\{2_n\}) = E_B[\Lambda_{12}(\{2_n\}\}]B)$$
 (1C)

for Gaussian background B is the expectation

$$I = E_{B} \left[exp \begin{cases} T & T \\ f & A(t)B(t)dt - c \int_{0}^{T} B^{2}(t)dt \end{cases} \right]$$

where

$$c = M/N$$

and

$$A(t) = \frac{2}{N} \sum_{n=1}^{M} (Z_n(t) - S_n(t))$$
 (20)

B(t) may be expanded in its Karhunen-Loeve (K-L) representation provided the mean square value

$$E(B^{2}(t)) < \infty$$
 (3C)

and thus the integral equation

$$\lambda_{i}\phi_{i}(t) = \int_{0}^{T} R_{B}(t,u) \phi_{i}(u) du \qquad 0 \leq t \leq T \qquad (4C)$$

can be solved. $\lambda_{\dot{1}}$ are called the eigenvalues of the equation and $\varphi_{\dot{1}}(t)$ are called the eigenfunctions.

Using Mercer's theorem, the correlation function of the background may be expressed as

$$R_{B}(t,u) = \sum_{k}^{\infty} \lambda_{k} \phi_{k}(t) \phi_{k}^{*}(u)$$
 (5C)

The K-L expansion of B(t) is

$$B(t) = \sum_{k=0}^{\infty} b_{k} \zeta_{k}(t)$$

with

$$b_{k} = \int_{0}^{T} B(\tau) \phi_{k}^{*}(\tau) d\tau \qquad 0 \le t \le T$$
 (60)

where the coefficients, \mathbf{b}_k , are uncorrelated and because B(t) is Gaussian the coefficients are statistically independent with means μ_k and variances λ_k .

Also expanding $\Lambda(t)$,

$$A(t) = \sum_{k}^{\infty} a_{k} \phi_{k}(t)$$

with

$$a_{k} = \int_{0}^{T} A(t) \phi_{k}^{*}(t) dt, \qquad 0 \le t \le T.$$
 (7C)

Rewriting in terms of the expansions

$$I = E_{B} \left[exp \begin{cases} \int_{0}^{T} \left(\frac{1}{2} \sum_{n=1}^{\infty} \sum_{k=1}^{\infty} a_{n} b_{k}^{*} + \frac{1}{2} \sum_{n=1}^{\infty} \sum_{k=1}^{\infty} a_{n}^{*} b_{k} \right) \phi_{k}(t) \phi_{n}^{*}(t) dt \right] - c \int_{0}^{T} \sum_{n=1}^{\infty} \sum_{k=1}^{\infty} b_{n} b_{k}^{*} \phi_{n}(t) \phi_{n}^{*}(t) dt \right].$$
(8C)

Using

$$\delta_{xy} = \int_{0}^{T} \phi_{x}(t) \phi_{y}^{*}(t) dt$$

$$I = E_{bn} \left[exp \left\{ \sum_{n=1}^{\infty} \left(\frac{1}{2} a_n b_n^{\star} + \frac{1}{2} a_n^{\star} b_n \right) - e \sum_{n=1}^{\infty} b_n b_n^{\star} \right\} \right]$$
(9C)

Using the independence of the coefficients

$$I = \prod_{n=0}^{\infty} E_{bn} \exp \left\{ \frac{1}{2} a_n b_n^* + \frac{1}{2} a_n^* b_n - c \left| b_n \right|^2 \right\}$$

$$= \sum_{n=0}^{\infty} \exp \left\{ \frac{\left| a_n \right|^2}{4c} \right\} E_{bn} \exp \left\{ -c \left| b_n - \frac{a_n}{2c} \right|^2 \right\}$$
(10C)

For Gaussian random variables

$$E(wxx^*) = \frac{\exp\left[w \cdot m_x m_x^*/(1-2w\lambda)\right]}{1-2w\lambda}$$

where

$$m_{\mathbf{x}} = \mathbf{E}(\mathbf{x})$$

$$\lambda = \mathbf{E}[(\mathbf{x} - \mathbf{m}_{\mathbf{x}})(\mathbf{x}^* - \mathbf{m}_{\mathbf{x}}^*)]$$
(110)

and w is a constant.

Let
$$x = b_n - a_n/2c$$

and
$$m_{\chi} = \mu_{\eta} - a_{\eta}/2c$$

then

$$E_{bn} \exp \left\{ -c \left| b_n - \frac{a_n}{2c} \right|^2 \right\} = \frac{\exp \left\{ -c \left| \mu_n - a_n / 2c \right|^2 (1 + 2c\lambda_n) \right\}}{(1 + 2c\lambda_n)}$$
 (12c)

Substituting (12C) into (10C), and separating terms

$$I = \prod_{n=1}^{B} (1 + 2c\lambda_{n})^{-1} \exp\left\{\frac{1}{2} \sum_{n=1}^{\infty} \frac{|a_{n}|^{2} \lambda_{n}}{1 + 2c\lambda_{n}}\right\}$$

$$\cdot \exp\left\{-\frac{c}{2} \sum_{n=1}^{\infty} \frac{\mu_{n}^{*}(\mu_{n} - a_{n}/c)}{1 + 2c\lambda_{n}}\right\}$$

$$\cdot \exp\left\{-\frac{c}{2} \sum_{n=1}^{\infty} \frac{\mu_{n}(\mu_{n}^{*} - a_{n}^{*}/c)}{1 + 2c\lambda_{n}}\right\}$$
(13C)

This result can be written in closed form by using the definition $\Lambda_n = (a, \phi_n)$

$$\frac{1}{2} \sum_{n}^{\infty} \frac{|a_{n}|^{2} \lambda_{n}}{1 + 2c\lambda_{n}} = \frac{N}{4M} \int_{0}^{T} dt_{1} \Lambda(t_{1}) \int_{0}^{T} dt_{2} \Lambda(t_{2}) q(t_{1}, t_{2})$$
 (14C)

where

$$q(t_1, t_2) = \frac{2M}{N} \sum_{n=1}^{\infty} \frac{\lambda_n}{1 + 2c\lambda_n} \phi_n(t_1) \phi_n^*(t_2)$$
 (150)

is such that operating on $q(t_1,t_2)$ with

$$\int_{0}^{T} \int_{0}^{T} \left[R_{B}(t_{1}, t_{2}) + \frac{1}{\sqrt{M}} R_{N}(T_{1}, t_{2}) \right] (\cdot) dt_{1} dt_{2}$$
 (15C)

and using Mercer's theorem yieles

$$\frac{N}{2M} q(t_1, t_2) + \int_{0}^{T} R_B(T_1, t_3) q(t_3, t_1) dt_3 = R_B(T_1, t_2)$$
 (170)

Likewise,

$$-\frac{c}{2} \sum_{n}^{\infty} \frac{\nu_{n}^{*} (\nu_{n}^{-a} - a_{n}^{\prime} c) + \nu_{n} (\nu_{n}^{*} - a_{n}^{*} / c)}{1 + 2c\lambda_{n}}$$

$$= \frac{1}{2} \int_{0}^{1} \left(m_{b}(t) - \frac{\Lambda(t)}{c} \right) h(t) dt \qquad (18C)$$

where

$$h(t) = \sum_{n=1}^{\infty} \frac{2\mu_{n}}{1 + 2c\lambda_{n}} \phi_{n}(t)$$
 (190)

such that

$$Nh(t) + 2M \int_{0}^{T} R_{B}(t_{3}, t) h(t_{3}) dt_{3} = Mm_{b}(t)$$
 (20C)

Taking the logarithm of both sides of equation (13C) and using the equations above

$$\begin{array}{lll}
\mathfrak{L}_{n} & 1 &= & \sum\limits_{n}^{\infty} \left(1 + 2c\lambda_{n}\right)^{-1} \\
& &+ & \frac{N}{4M} \int_{0}^{T} dt_{1} \left\{ \frac{2}{N} \sum\limits_{k=1}^{M} \left(Z_{k}(t_{1}) - S_{k}(t_{1})\right) \right\} \\
& &+ \int_{0}^{T} dt_{2} \left\{ \frac{2}{N} \sum\limits_{m=1}^{M} \left(Z_{m}(t_{2}) - S_{m}(t_{2})\right) \right\} q(t_{1}, t_{2}) \\
& &+ \int_{0}^{T} \left[m_{b}(t) - \frac{N}{M} \left\{ \frac{2}{N} \sum\limits_{k=1}^{M} \left(Z_{k}(t) - S_{k}(t)\right) \right\} \right] h(t) dt \\
& + \frac{1}{NM} \int_{0}^{T} \int_{0}^{T} \left[\sum\limits_{k=1}^{M} Z_{k}(t_{1}) \sum\limits_{m=1}^{M} Z_{m}(t_{2}) - \sum\limits_{k=1}^{M} S_{k}(t_{1}) \sum\limits_{m=1}^{M} Z_{m}(t_{2}) - \sum\limits_{k=1}^{M} S_{k}(t_{1}) \sum\limits_{m=1}^{M} Z_{m}(t_{2}) - \sum\limits_{m=1}^{M} S_{m}(t_{2}) \sum\limits_{m=1}^{M} Z_{m}(t_{2}) - \sum\limits_{k=1}^{M} S_{k}(t_{1}) \sum\limits_{m=1}^{M} S_{m}(t_{2}) \right] q(t_{1}, t_{2}) dt_{1} dt_{2} \\
& + \int_{0}^{T} \sum\limits_{m=1}^{M} (T_{k}) \sum\limits_{k=1}^{M} (T_{k}) dt_{1} dt_{2} \\
& + \frac{2}{M} \int_{0}^{T} \sum\limits_{k=1}^{M} (T_{k}) \sum\limits_{k=1}^{M} (T_{k}) \sum\limits_{k=1}^{M} (T_{k}) \sum\limits_{k=1}^{M} (T_{k}) \sum\limits_{k=1}^{M} (T_{k}) dt_{1} dt_{2} \\
& + \frac{2}{M} \int_{0}^{T} \sum\limits_{k=1}^{M} (T_{k}) \sum\limits_{k=1}^{M} (T_{k}) \sum\limits_{k=1}^{M} (T_{k}) \sum\limits_{k=1}^{M} (T_{k}) \sum\limits_{k=1}^{M} (T_{k}) dt_{1} dt_{2} \\
& + \frac{2}{M} \int_{0}^{T} \sum\limits_{k=1}^{M} (T_{k}) \sum\limits_{k=1}^{M} (T_{k}) \sum\limits_{k=1}^{M} (T_{k}) \sum\limits_{k=1}^{M} (T_{k}) \sum\limits_{k=1}^{M} (T_{k}) \sum\limits_{k=1}^{M} (T_{k}) dt_{1} dt_{2} \\
& + \frac{2}{M} \int_{0}^{T} \sum\limits_{k=1}^{M} (T_{k}) \sum\limits_{k=1}^{M} (T_{k}) \sum\limits_{k=1}^{M} (T_{k}) \sum\limits_{k=1}^{M} (T_{k}) \sum\limits_{k=1}^{M} (T_{k}) dt_{1} dt_{2} \\
& + \frac{2}{M} \int_{0}^{T} \sum\limits_{k=1}^{M} (T_{k}) \sum\limits_{k=1}^{M} (T_{k}) \sum\limits_{k=1}^{M} (T_{k}) \sum\limits_{k=1}^{M} (T_{k}) dt_{1} dt_{2} \\
& + \frac{2}{M} \int_{0}^{T} \sum\limits_{k=1}^{M} (T_{k}) \sum\limits_{k=1}^{M} (T_{k}) \sum\limits_{k=1}^{M} (T_{k}) \sum\limits_{k=1}^{M} (T_{k}) dt_{1} dt_{2} \\
& + \frac{2}{M} \int_{0}^{T} \sum\limits_{k=1}^{M} (T_{k}) \sum\limits_{k=1}^{M} (T_{k}) \sum\limits_{k=1}^{M} (T_{k}) dt_{1} dt_{2} \\
& + \frac{2}{M} \int_{0}^{T} \sum\limits_{k=1}^{M} (T_{k}) \sum\limits_{k=1}^{M} (T_{k}) \sum\limits_{k=1}^{M} (T_{k}) dt_{1} dt_{2} \\
& + \frac{2}{M} \int_{0}^{T} \sum\limits_{k=1}^{M} (T_{k}) \sum\limits_{k=1}^{M} (T_{k}) dt_{1} dt_{2} \\
& + \frac{2}{M} \int_{0}^{T} \sum\limits_{k=1}^{M} (T_{k}) dt_{1} dt_{2} dt_{1} dt_{2} dt_{1} dt_{2} dt_{1} dt_{2} dt_{1} dt_{2} dt_{2} dt_{2} dt_{2} dt_{2} dt_{2} dt_{2} dt_{2$$

The likelihood ratio desired is

$$\Lambda_{12}(\{Z_n\}) = \exp\left\{\frac{2}{N} \sum_{n=1}^{M} \int_{0}^{T} Z_n(t) S_n(t) dt - \frac{1}{N} \sum_{n=1}^{M} \int_{0}^{T} S_n^2(t) dt\right\} \cdot I.$$
 (22C)

Taking the logarithm of both sides and substituting the expression for $\mathbb{A}_n(I)$,

$$\begin{split} & \ln \Lambda_{12}(\{Z_n\}) = \frac{2}{N} \sum_{n=1}^{M} \int_{0}^{T} Z_n(t) S_n(t) dt \\ & - \frac{1}{N} \sum_{n=1}^{M} \int_{0}^{T} S_n^2(t) dt \\ & + \sum_{k=1}^{\infty} \ln (1 + \frac{2M}{N} \lambda_k)^{-1} \\ & + \frac{1}{NM} \int_{0}^{T} \int_{0}^{T} \left\{ \sum_{k=1}^{M} Z_k(t_1) \sum_{m=1}^{M} Z_m(t_2) - \sum_{m=1}^{M} S_m(t_2) \sum_{k=1}^{M} Z_{k}(t_1) \right\} \\ & - \sum_{k=1}^{M} S_k(t_1) \sum_{m=1}^{M} Z_m(t_2) - \sum_{m=1}^{M} S_m(t_2) \sum_{k=1}^{M} Z_{k}(t_1) \\ & + \sum_{k=1}^{M} S_k(t_1) \sum_{m=1}^{M} S_m(t_2) \right\} q(t_1, t_2) dt_1 dt_2 \\ & - \int_{0}^{T} m_b(t) h(t) dt + \frac{2}{M} \int_{0}^{T} \sum_{k=1}^{M} (Z_k(t) - S_k(t)) h(t) dt \end{split}$$

(23C)

APPENDIX D

KARHUNEN-LOEVE REPRESENTATION OF EQUIVALENT SIGNAL-TO-NOISE RATIO

The variance of the threshold statistic

$$W = \sum_{n=1}^{M} \int_{0}^{T} Z_{n}(t) \bar{q}(t) dt$$

where

$$\bar{q}(t) = \frac{2}{N} \left[\bar{f}(t) - \int_{0}^{T} \bar{f}(u) q(u,t) du \right]$$
 (1D)

under hypotheses H_o and H_1 is

$$VAR = \frac{NM}{z} \int_{0}^{T} \tilde{q}^{2}(t) dt + M^{2} \int dt \cdot \tilde{q}(t) \int du \tilde{q}(u) R_{y}(t, u)$$
 (2D)

Substituting (1D) into the first term of the variance expression

$$\frac{NM}{z} \int_{0}^{T} \bar{q}^{2}(t) dt = \frac{2M}{N} \int_{0}^{T} dt \left[\bar{f}^{2}(t) - 2\bar{f}(t) \int_{0}^{T} \bar{f}(u) q(t, t) du \right]$$

$$+ \int_{0}^{T} \int_{0}^{T} \bar{f}(u) \bar{f}(v) q(t, u) q(t, v) du dv$$
(3D)

Using the Karhunen-Loeve representation

$$\frac{NM}{z} \int_{0}^{T} \tilde{q}^{2}(t) dt = \frac{2M}{N} \left[\sum_{k=1}^{\infty} \tilde{f}_{k} \tilde{f}_{k}^{*} \int_{0}^{T} \phi_{k}(t) \phi_{k}^{*}(t) dt \right]$$

$$-2 \sum_{k=1}^{\infty} \sum_{k=1}^{\infty} \tilde{f}_{k} \tilde{f}_{k}^{*} \frac{2M}{1 + \frac{2M}{N}} \frac{\lambda_{n}}{\lambda_{n}} \int_{0}^{T} \phi_{k}(t) \phi_{m}^{*}(t) dt \int_{0}^{T} c_{m}(u) \phi_{k}^{*}(u) du$$

$$+ \sum_{k} \sum_{m} \sum_{n} \sum_{n} f_{k} f_{k}^{*} \frac{\frac{2m}{N} \lambda_{m}}{1 + \frac{2m}{N} \lambda_{m}} \frac{\frac{2m}{N} \lambda_{n}}{1 + \frac{2m}{N} \lambda_{n}} \int_{0}^{T} \phi_{m}(t) \phi_{n}^{*}(t) dt$$

$$\int_{\mathbf{0}}^{\mathbf{T}} \phi_{k}(\mathbf{u}) \phi_{n}^{*}(\mathbf{u}) d\mathbf{t} \int_{\mathbf{0}}^{\mathbf{T}} \phi_{n}(\mathbf{v}) \phi_{k}^{*}(\mathbf{v}) d\mathbf{t}$$
(4D)

Using

$$\delta_{nk} = \int_{0}^{T} \phi_{n}(t) \phi_{k}^{*}(t) dt$$

then

$$\frac{NM}{z} \int_{0}^{T} \tilde{q}^{2}(t) dt = \frac{2M}{N} \sum_{k}^{\infty} \tilde{f}_{k} \tilde{f}_{k}^{*} \left(1 - 2 \frac{\frac{2M}{N} \lambda_{k}}{1 + \frac{2M}{N} \lambda_{k}} + \left(\frac{\frac{2M}{N} \lambda_{k}}{1 + \frac{2M}{N} \lambda_{k}}\right)^{2}\right)$$

$$= \frac{2M}{N} \sum_{k}^{\infty} \tilde{f}_{k} \tilde{f}_{k}^{*} \left(\frac{1}{1 + \frac{2M}{N} \lambda_{k}}\right)^{2}. \tag{5D}$$

The second term may be expressed in a similar manner to give

$$VAR = \frac{2M}{N} \sum_{k=1}^{\infty} \tilde{f}_{k} \tilde{f}_{k}^{*} \left(\frac{1}{1 + \frac{2M}{N} \lambda_{k}} \right)^{2} + \left(\frac{2M}{N} \right)^{2} \sum_{k=1}^{\infty} \tilde{f}_{k} \tilde{f}_{k}^{*} \left(\frac{1}{1 + \frac{2M}{N} \lambda_{k}} \right)^{2} \lambda_{k}$$

$$VAR = \frac{2M}{N} \sum_{k}^{\infty} \bar{f}_{k} \bar{f}_{k}^{*} \frac{1}{1 + \frac{2M}{N} \lambda_{k}}. \qquad (6D)$$

The equivalent signal to noise ratio in Part II is

$$d = \frac{\sum_{n=1}^{M} \tilde{a}_n \int_{0}^{T} \tilde{f}(t)\tilde{q}(t)dt}{\sqrt{VAR}}$$
 (7D)

The K-L representation of the numerator is

$$\sum_{n=1}^{M} \tilde{a}_{n} \int_{0}^{T} \tilde{f}(t) \tilde{q}(t) dt = \sum_{n=1}^{M} \tilde{q}_{n} \left[\frac{2}{N} \sum_{k=1}^{\infty} |\tilde{f}_{k}|^{2} / \left(1 + \frac{2M}{N} \lambda_{k} \right) \right]. \quad (8D)$$

d may now be represented by dividing (8D) by the square root of (6D)

$$d = \frac{\sum_{N=1}^{2} \frac{M}{N} a_{N} \frac{\sum_{k} |\tilde{f}_{k}|^{2} / (1 + \frac{2M}{N} \lambda_{k})}{\left[\frac{2M}{N} \sum_{k} |\tilde{f}_{k}|^{2} / (1 + \frac{2M}{N} \lambda_{k})\right]^{\frac{1}{2}}}$$

$$d = \frac{\sum_{N=1}^{2} \frac{M}{N} a_{N} \sqrt{N} \left[\sum_{k} |\tilde{f}_{k}|^{2} / (1 + \frac{2M}{N} \lambda_{k})\right]^{\frac{1}{2}}}{\sqrt{N/2}}.$$
 (90)

APPENDIX E

NON-ZERO TERMS FOR NUMERICAL INTEGRATION

Equation (5.5) for the radiance function for a point target may be written as

$$1 = \frac{j^2}{\pi^2} \int e^{-x^2 - y^2} e^{-j(P_X + Q_y)} \frac{\sin \alpha x}{x} \frac{\sin \beta y}{y}$$

$$\cdot e^{-jn\phi} J_n\left(c\sqrt{x^2 + y^2}\right) dxdy$$

where

$$\alpha = w/\sqrt{2}o$$

$$\beta = h/\sqrt{2}o$$

$$c = \sqrt{2}\rho/\sigma$$

$$p = \frac{\sqrt{2}}{\sigma} (r_x - x_0)$$

$$Q = \frac{\sqrt{2}}{6} (r_y \cdot y_0)$$

and

$$\phi = \tan^{-1} y/x \tag{1E}$$

Let E(x,y) be that part of the integral that is even about both planar axis which is

$$E(x,y) = \frac{j^n}{\pi^2} e^{-x^2 - y^2} \frac{\sin \alpha x}{x} \frac{\sin \beta y}{y} J_n \left(c \sqrt{x^2 + y^2} \right).$$
 (2E)

Using Euler's relation the integral becomes

 $I = \int E(x,y) \left[\cos(Px+Qy) + j\sin(Px+Qy)\right] \left[\cos n\phi - j\sin n\phi\right] dxdy \qquad (3E)$ Using trigonometric identities

$$1 = \int E(x,y) \sum_{i=1}^{\infty} T_i dxdy$$

where

$$T_{1} = \cos Px \cos Qy \cos n\phi$$

$$T_{2} = -\sin Px \sin Qy \cos n\phi$$

$$T_{3} = -\sin Px \cos Qy \sin n\phi$$

$$T_{4} = -\cos Px \sin Qy \sin n\phi$$

$$T_{5} = -j \sin Px \cos Qy \cos n\phi$$

$$T_{6} = -j \cos Px \sin Qy \cos n\phi$$

$$T_{7} = -j \cos Px \cos Qy \sin n\phi$$

$$T_{8} = +j \sin Px \sin Qy \sin n\phi$$
(4E)

To save computation time it is necessary to determine which of the eight integrals are always zero and for which is they are zero.

Observe that

$$\cos n\phi = R_e e^{jn\phi} = R_e(e^{j\phi})^n$$

$$= R_e(\cos\phi + j \sin\phi)^n$$

$$= R_e \left(\frac{x}{\sqrt{x^2 + y^2}} + j \frac{y}{\sqrt{x^2 + y^2}}\right)^n$$

$$= \left(\frac{1}{x^2 + y^2}\right)^{n/2} R_e(x + jy)^n. \tag{SE}$$

Likewise

$$\sin n\phi = \left(\frac{1}{x^2 + y^2}\right)^{n/2} I_m(x+jy)^n . \tag{6E}$$

The first term in brackets is always even and may be made part of E(x,y). The second term in brackets may be expanded by the binomial expansion with the general term

$$(x)^{n-k} (jy)^k . (7E)$$

The first integral of (4E) becomes

$$I_1 = \int E(x,y)\cos Px \cos Qy R_e(x+jy)^n dxdy$$
 (8E)

For I_1 to be non-zero, $Re(x+jy)^n$ must be even in x and even in y. From (7E) this is true only of even n; for odd n, I_1 will always be zero.

Examining the other seven integrals yields

$$I = \begin{cases} \int E(x,y) [\cos Px \cos Qy \cos n\phi] \\ + j \sin Px \sin Qy \sin n\phi] dxdy & n \text{ even} \end{cases}$$

$$\int E(x,y) [\cos Px \sin Qy \sin n\phi] \\ + j \sin Px \cos Qy \cos n\phi] dxdy & n \text{ odd} \end{cases}$$

$$(9E)$$

Observing that symmetry about the x- and y-axis is necessary for this integral to be non-zero and integrating only when the integral is non-zero permits one to integrate only over the first quadrant.

APPENDIX F

SUMMATION CHECK

The signal for a point target was shown in Part II to be

$$H(\omega_{o}t) = \int N'(r)\tau(r-\rho(t)) d^{2}r . \qquad (1F)$$

Assuming an infinite image plane, $H(\omega_0^-t)$ may be written in terms of the Fourier transform as

$$H(\omega_0 t) = \int N'(k) \tau^{*}(k) e^{-j2\pi k \cdot \rho(t)} d^2k$$
 (2F)

For a rectangular detector (2F) may be written as

$$H(\omega_{o}t) = \int e^{-2\pi^{2}o^{2}(k_{x}^{2}+k_{y}^{2})} \frac{\sin\pi wkx}{\pi k_{x}} \frac{\sin\pi ky}{\pi k_{y}}$$

$$\cdot \int e^{j2\pi k \cdot (\varrho(t) - r_{o})} \cdot e^{-j2\pi(k_{x}x_{o} + k_{y}y_{o})} d^{2}k \quad (3F)$$

Let

$$x = \sqrt{2} \circ \pi kx$$

$$y = \sqrt{2} \circ \pi ky$$

and noting the integral may be separated into two parts

$$H(\omega_{o}t) = \frac{1}{\pi} \int e^{-x^{2}} \sin \frac{wx}{\sqrt{2}\sigma} e^{j\frac{\sqrt{2}}{\sigma} x(\rho\cos\omega_{o}t - (r_{x}-x_{o}))} dx$$

$$\cdot \frac{1}{\pi} \int e^{-y^{2}} \frac{\sin \frac{hy}{\sqrt{2}\sigma}}{y} e^{j\frac{\sqrt{2}}{\sigma} y(\rho\sin\omega_{o}t - (r_{y}-y_{o}))} dy . \tag{4F}$$

Since the integration is $-\infty$ to $+\infty$ only even integrands will be non-zero. Rewriting (4F)

$$H(\omega_0 t) = \frac{2}{T} \int_0^\infty e^{-x^2} \frac{\sin \frac{wx}{\sqrt{2}\sigma}}{x} \cos \left(\frac{\sqrt{2}}{\sigma} x [\rho \cos \omega_0 t - (r_x - x_0)]\right) dx$$

$$\cdot \frac{2}{\pi} \int_{0}^{\infty} e^{-y^{2}} \frac{\sin \frac{hy}{\sqrt{2}\sigma}}{y} \cos \left(\frac{\sqrt{2}}{\sigma} y \left[\rho \sin \omega_{o} t - (r_{y} - y_{o})\right]\right) dy.$$
(5F)

The general form of the integral is

$$R = R(B) = \frac{2}{\pi} \int_{0}^{\infty} e^{-u^2} \frac{\sin By}{y} \cos(Cu) du.$$
 (6F)

Differentiating with respect to B

$$\frac{dR(B)}{dB} = \frac{2}{\pi} \int_{0}^{\infty} e^{-u^{2}} \cos By \cos Cu \ du.$$
 (7F)

But the solution to (7F) is known [2]:

$$\frac{dR(B)}{dB} = \frac{1}{2\sqrt{\pi}} \left\{ e^{-\frac{(B-C)^2}{4}} + e^{-\frac{(B+C)^2}{4}} \right\}.$$
 (8F)

Since the constant of integration R(0) = 0,

$$R(B) = \frac{1}{2\sqrt{\pi}} \left[\int_{0}^{\infty} e^{-\frac{(v-c)^{2}}{4}} dv + \int_{0}^{\infty} e^{-\frac{(w+c)^{2}}{4}} dw \right]$$
 (9F)

Let

$$P = \frac{v-c}{2}$$

and

$$Q = \frac{W+C}{2}$$

then

$$R(B) = \frac{1}{2} \left\{ \frac{2}{\sqrt{\pi}} \int_{-C/2}^{\frac{B-C}{2}} e^{-P^2} dP + \frac{2}{\sqrt{\pi}} \int_{-C/2}^{\frac{B+C}{2}} e^{-Q^2} dQ \right\}$$

$$= \frac{1}{2} \left\{ Erf \left(\frac{B-C}{2} \right) + Erf \left(\frac{B+C}{2} \right) \right\}$$
(10F)

where

$$Erf(z) = \frac{2}{\sqrt{\pi}} \int_{0}^{z} e^{-y^{2}} dy$$
 (11F)

Substituting the result (10F) into (5F)

$$H(\omega_{o}t) = \frac{1}{4} \left\{ \text{Erf} \left(\frac{\sqrt{2}}{T} \left[\rho \cos \omega_{o}t - r_{x} + x_{o} \right] + \frac{1}{\sqrt{2\sigma}} w \right) - \text{Erf} \left(\frac{\sqrt{2}}{\sigma} \left[\rho \cos \omega_{o}t - r_{x} + x_{o} \right] - \frac{1}{\sqrt{2\sigma}} w \right) + \text{Erf} \left(\frac{\sqrt{2}}{\sigma} \left[\rho \sin \omega_{o}t - r_{y} + y_{o} \right] + \frac{1}{\sqrt{2\sigma}} h \right) - \text{Erf} \left(\frac{\sqrt{2}}{\sigma} \left[\rho \sin \omega_{o}t - r_{y} + y_{o} \right] - \frac{1}{\sqrt{2\sigma}} \right) \right\}. \quad (12F)$$

The mean square voltage of the background may also be solved in a like manner.

Frem Part V, the mean square voltage was shown to be

$$\tilde{\mathbf{e}}_{\mathbf{B}}^{2} = \int |\tau(\underline{k})|^{2} |F_{o}(\underline{k})|^{2} |W_{\mathbf{B}}(\underline{k})|^{2} |W_{\mathbf{B}}(\underline{k})|^{2}$$
(13F)

which for a rectangular detector becomes

$$\tilde{c}_{B}^{2} = \sigma_{\beta} \cdot 2\alpha \int \frac{\sin^{2}(\pi k_{x} W)}{\pi^{2} k_{x}^{2}} \frac{e^{-4\pi^{2}\sigma^{2} k_{x}^{2}}}{\sigma^{2} + (2\pi k_{x})^{2}} dk_{x}$$

$$+\sigma_{B} + 2\beta \int \frac{\sin^{2}(\pi k_{y}h)}{\pi^{2}k_{y}^{2}} = \frac{e^{-4\pi^{2}\sigma^{2}k_{y}^{2}}}{e^{2} + (2\pi k_{y})^{2}} dk_{y}$$
 (14F)

Letting

$$t = 2\pi k$$

the general form of the integral to be solved is

$$R(\mu) = \frac{1}{\pi} \int_{0}^{\infty} \frac{\sin^{2}(\mu \tau/2)}{t^{2}/4} \frac{e^{-\sigma^{2}t^{2}}}{\sigma^{2} + t^{2}}$$
 (15F)

Taking the first and second derivatives one has

$$\frac{dR(u)}{du} = \frac{1}{\pi} \int_{0}^{\infty} \frac{\sin(2\mu t)}{t} = \frac{e^{-\sigma^{2}t^{2}}}{\alpha^{2}+t^{2}} dt$$
 (16F)

and

$$\frac{d^{2}R(u)}{du^{2}} = \frac{2}{\pi} \int_{0}^{\infty} \frac{\cos(2ut)}{\alpha^{2} + t^{2}} e^{-\sigma^{2}t^{2}}.$$
 (17F)

The solution to the second derivative is known [2]:

$$\frac{d^{2}R(\mu)}{du^{2}} = \frac{e^{\sigma^{2}\alpha^{2}}}{2\alpha} \left[e^{-\alpha\mu} \text{ Erfc } (\alpha\sigma - \frac{\mu}{2\sigma}) + e^{\alpha\mu} \text{ Erfc } (\alpha\sigma + \frac{\mu}{2\sigma}) \right]$$
(18F)

where

$$\operatorname{Erfc}(z) = \frac{2}{\sqrt{\pi}} \int_{z}^{\infty} e^{-y^{2}} dy$$
 (19F)

is the complimentary error function.

With lengthy and tedious calculations it can be shown by using the relation

$$R(\mu) = \int_0^{\mu} dv \left[\int_0^{v} F(w) dw \right]$$
 (20F)

with initial conditions equal to zero and

$$F(w) = \frac{d^2R(u)}{du^2}$$
 (21F)

that

$$R(\mu) = \frac{\mu}{\alpha^2} \left[1 - \operatorname{Erfc} \left(\frac{\mu}{2\sigma} \right) - \frac{2\sigma}{\sqrt{\pi}u} \left(1 - e^{-(\mu/2\sigma)^2} \right) \right]$$

$$+ \frac{e^{\sigma^2 \alpha^2}}{2\alpha^3} \left[e^{\alpha \mu} \operatorname{Erfc} \left(\alpha \sigma + \frac{\mu}{2\sigma} \right) + e^{-\alpha \mu} \operatorname{Erfc} \left(\alpha \sigma - \frac{\mu}{2\sigma} \right) \right]$$

$$- 2 \operatorname{Erfc} \left(\alpha \sigma \right) \right] \qquad (22F)$$

The mean square voltage thus becomes

$$\tilde{e}_{B}^{2} = 4\alpha\beta\sigma_{B}^{2} R(w;\alpha)R(h;\beta) . \qquad (23F)$$

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APPENDIX H

COMPUTER PROGRAM TO CALCULATE THE FOURIER CORFFICIENTS OF THE BACKGROUND CORRELATION FUNCTION, A POINT TARGET AND AN AVERAGED TARGET OVER A SAUSSIAL POINTING ERROR FOR A NUTATING OPTICAL SYSTEM WITH A RECTANGULAR DETECTOR.

AA = BACKGROUND CORRELATION FUNCTION COEFF(ZERO GRDER)
A| FA = BACKGROUND CORRELATION COEFFICIENTS (VECTOR)
FNO = SIGNAL FOURIER COEFFICIENT (ZERO ORDER)
FN = SIGNAL FOURIER COEFFICIENTS (VECTOR)
FAVOD = AVERAGED SIGNAL COEFF (ZERO ORDER)
FAVOD = AVERAGED SIGNAL COEFF (ZERO ORDER)
FAVOD = AVERAGED SIGNAL FOURIER COEFF (VECTOR)
RHO = NUTATION RADIUS
SIG = SIANDARD DEVIATION OF BLUR CIRCLE
WI = WIDTH OF DETECTOR (X-DIRECTION)
H = HEIGHT OF DETECTOR (Y-DIRECTION)
AI = INVERSE COER LENGTH OF BACKGROUND (X-DIRECTION)
BF = INVERSE COER LENGTH OF BACKGROUND (Y-DIRECTION)
PO = SIANDARD DEVISTION OF PAINTING ERROR
XO = DETECTOR COORDINATE X-DIRECTION
YO = DETECTOR COORDINATE X-DIRECTION
RX = POINT TARGET COORDINATE X-DIRECTION
RY = POINT TARGET COORDINATE X-DIRECTION
RY = POINT TARGET COCRDINATE X-DIRECTION
RY = POINT TARGET COEFFFICIENT DESIRED
NUM = NUMBER OF LAST COFFFICIENT DESIRED
NUM = NUMBER OF LAST COFFFICIENT DESIRED
IAVE=0 "SIGNAL" WILL CALCULATE POINT TARGET COEFF

IMPLICIT RÉALMO (A-H.D-2)
DIMENSION ALHA(120)
COMMPLÉXAIG FA(120).FNO.FAVG(120).FAVGO
COMMON /DIPRM/ RHO.SIG.WI.H.AL.DE.RX.RY.XO.YO
COMMON /AVPRM/ NUM.IAVÉ.PD
COMMON /AVPRM/ NUM.IAVÉ.PD
COMMON /ANGL / RANGL
READ 2.RHJ.SIG.WI.H.AL.BE
READ 2.XO.YO.R.ANGL
READ 3.NS.NUM.IAVE.PD
RANGLEO.0174500.ANGL
RX=R-DCCS(KANGL)
RX=R-DCCS(KANGL)
RX=R-DCCS(KANGL)
RANGLEDATAN2(RY-YO.RX-XO)
CALL SUMOK
CALL BKGEN(ALFA.AA)
CALL SIGNAL(FN.FNO)
WRITE (9) RHO.SIG.WI.H.AL.BE.RX.RY.XO.YO.PD
IAVE=1
IF(IAVE.ED.1) RX=0.0D0
IF(IAVE.ED.1) RX=0.0D0
CALL SIGNAL(FAVG.FAVG.)
WRITE (9) FAVG.FAVG.)
WRITE (9) FAVG.FAVG.
FORMAT(1015)
FORMAT(315.1015.0)
FORMAT(315.1015.0)
FORMAT(315.1015.0)
FORMAT(315.1015.0)
FORMAT(8F10.44)
STOP
END

123

```
SUBROUTINE BKCRN(VP.V)
SUBROUTINE BEGRN COMPUTES THE COEFFICIENTS OF THE BACKGROUND CORRELATION FUNCTION UP TO 120 ORDERS. HIGHER ORDERS MAY EASTLY BE COMPUTED BY ONLY CHANGING THE DIMENSION STATEMENTS.
                 ZP = VALUES OF RADII FOR DIFFERENT ANNULI
Z = WORKING VECTOR FOR RADII
MAXR = MAXIMUM RADIUS OF OUTER ANNULUS
NIZI = DIMENSION OF ZP
VP = WORKING VECTOR IN FIRST PART AND RETURN VECTOR
IN THE SECOND PART
V = ZERO ORDER COEFFICIENT
              IMPLICIT REAL"8 (A-H, C-Z)
DIMENSION VP(1).DN(50.120)
DIMENSION DI(50).ZP(20).Z(50)
DIMENSION P(74).W(24).TH(25)
COMMON /AVPRN/ NUM.HAVE.PD
COMMON /DTPRM/ RHO,SIG.WI,H,AL,BE,RX,RY,XO,YO
COMMON /BKPRM/ A.B.C.D
COMMON /NFIRST/ NFIRST
DATA MAXE/4/.NR/1/.NIZ1/20/
DATA ZP/0.070.1.200,Z.400.5.500.8.600.11.800.15.00.20.
100.25.000.30.070.40.070.50.070.60.070.70.000,80.000,
290.070.100.000.150.000.200.000.400.000/
NFIRST=0
PI = DARCOS(-1.00)
                 PI = DARCOS (-1.00)
                 CALCULATE CONSTANTS FOR INTEGRAL
                CCCC
                  SET ANNULI FOR SPECIFIC PARAMETERS
                DO 150 I=1.NIZ1
Z(I)=ZP(I)/E

I = 0
IF(Z(NIZ1).GT.MAXR)GD TC 210
DZ = P!/NR/E
MMM=50-NIZ1
DO 200 I=1.MMM
Z(I+NIZ1) = Z(NIZ1) + I*DZ
IF(Z(I+NIZ1).GT.MAXR)GO TD 2
CONTINUE
   150
      200 CONTINUE
MAX = NUMBER OF ANNULI
   210
                  MAX=1+V17.1-1
                 PRINT 4, RHO. $13, H, WI. AL, 85, A, B, C, D, E, DZ PRINT 4, (Z(I), I=1, MAX) DO 300 I=1, MAX
COMPUTE RADII, WEIGHTS AND ANGLES FOR EACH ANNULUS
```

```
CALL WAA(Z(I),Z(I+1),R,TH,W,NPQ)
C
              NFIRST=1
D1(1) = 0.00
D0 115 NN=1.NUM
DN(1,NN) = 0.00
              COMPUTE FUNCTION FOR EACH RADIUS
              DO 280 K*1, NPQ
CALL 3ESSL (R(K) "E, VP, VJN, NUM)
RV=VJN-32/DEXP(F(K) # \2)
EP=DEXP(R(K) * \*2)
              DO 120 NN=1.NUM
VP(NN)=VP(NN)#+2/=P
              REC=0.DO
CCCC
              COMPUTE FUNCTION FOR EACH ANGLE
              DO 270 J=1,NPO
REC=REC+FU(R(K)+DCOS(TH(J)),K(K)+DSIN(TH(J)))
   270
               V= O.DO
              DD 130 NN=1, NUM

DN(1,NN)=DN(1,NN)+REC&VP(NN)&W(K)

DI(1) = DI(1) + REC%RV*W(K)

CONTINUE
  130
280
300
               SUM VALUE FROM FACH ANNULUS
              \begin{array}{l} DG & 400 & I = 1. \text{MAX} \\ J & = \text{MAX} & + 1 & - 1 \\ V & = V & + DI & (J) \end{array}
               V=V*CONST
PRINT 3.N.V.DI(Mex)
              PRINT 3.N.V.DICTEX,
NBQ=0
PRINT 13.NBQ,V
PRINT 4.(DI(1),I=1.MAX)
DD 500 NP=1.NUM
VQM=0.D0
VP(NP)=0.D0
DU 600 I=1.MAX
I=MAX+1-1
               J=MAX+1-1
J=MAX+1-1
VP(NP)=VP(NP)+DN(J,NP)
IF(DN(J,NP).GT.VQM) VOM=CN(J,NP)
   600
               VP(NP)=VP(NP)*CONST
PRINT 13,NP,VP(NP)
PRINT 14,NP,VQM, DN(MAX,NP)
   500
               CONTINUE
               CALCULATE SUM FOR INTEGRATION CHECK
               SUM=V
DD 700 J=1,NUM
I=NUM+1-J
SUM=SUM+2.DO+VP(I)
              CONTINUE
PEINT 12, SUM
RETURN
   700
              RETURN

FORMAT (1015)

FORMAT (2110.0)

FORMAT (2110.1P2515.6)

FORMAT (168515.6)

FORMAT (5X.620.13)

FORMAT (5X.14.6X.620.13)

FORMAT (5X.14.6X.620.13,6X.620.13)

FORMAT (5(1PD20.12))
   1
2
3
   4
12
13
14
   16
               E ND
```

FUNCTION FU(X,Y)

FUNCTION FU IS AN AUXILIARY ROUTINE FOR BEGRN THAT COMPUTES THAT PART OF THE INTEGRAND NOT RADIALLY SYMMETRIC.

IMPLICIT REALMS (A-H, C-Z)
CCMMON /BKPRM/ A,B,C,D
P1=1.D0
IF(X.NE.O.DO) P1=DSIN(AMX)/(AMX)
IF(Y.NE.O.DO) P1=P1+CSIN(BMY)/(BMY)
FU=P1+42/(C+XM-2)/(C+YM-2)
RETURN
END

```
CCC
                  MAX = NUMBER OF ANNULI
                  MAX=I+NIZ1-1
GD TO 212
MAX=1
   210
   211
212
215
220
                  PRINT 4, KHO, SIG, WI, H, AL, BE, RX, RY, XO, YO, PX, QX
PRINT 4, (Z(I), I=1, MAX)
DC 300 I=1, MAX
CCCC
                   COMPUTE RADII, WEIGHTS AND ANGLES FOR EACH ANNULUS
                  CALL WAA(Z(I),Z(I+1),R,TH,W,NPQ)
NFIRST=1
NPR=NPQ
                  NPR=NPQ

DI(I) = 0.00

DG 115 NN=NS, NUM

DNJ(I,NN) = 0.00

IF(KKK.EQ.1) GC TO 300

DC 290 K=1,NPQ

RK2=R(K)$*2

IF(RK2.GT.174) GO TC 285

EP=DEXP(RK2)

IF(EP.GT.1.0 50) GC TO 285

CALL BESSL (P(K)+E,BN,VJN,NUM)

RV=VJN/FP

RECG=F(R(K).TH.0)
    115
                   RECG=F(R(K),TH,O)
                    COMPUTE FUNCTION FOR EACH RADIUS
                  RECO=RECO*RV*W(K)
PP=RECO
D1(I)*D1(I)*PP
DG 130 NN=NS, MJM
BN(NN) = BN(NN) / EP
REC(NN) = F(P(K), TH, NN)
REC(NN) = REC(NN) * BN(NN) * W(K)
PP=REC(NN)
DN(I+NN) = VN(I+NN) + PP
PP=REC(NN)
DNJ(I+NN) = UNJ(I+NN) + PP
CONTINUE
CONTINUE
CONTINUE
   130
280
300
                    CONTINUE
                    SUM VALUE FROM EACH ANNULUS
                    VZ = 0.00
V1 = 0.00
V2 = 0.00
                   V2 = 0.00

D0 400 I=1,MAX

J = MAX + 1 - I

If(DABS(DI(J)).GT.CABS(VZ)) VZ=DI(J)

IF(DI(J).LE.0.00) V1 = V1+DI(J)

IF(DI(J).GT.0.00) V2 = V2+DI(J)

RECO=DCMPLX(V1+V2,0.00)

RECO=CCASI
    400
                 RECURRECURCING.
NBO=0
PRINT 4.(DI(1).I=1,MAX)
PRINT 19,NAG,RECO
PRINT 14,NBG,VZ,DI(MAX)
DO 500 NP=NS,NJM
VQM = 0.DO
V1 = 0.DO
V1 = 0.CO
V2 = 0.CO
V21=0.DO
```

```
DO 600 I=1.MAX
                             J=MAX+1-I

1F(DN(J,NP).LT.O.DO) V1=V1+DN(J,NP)

IF(DNJ(J,NP).LT.O.DO) V1J=V1J+DNJ(J,NP)

IF(DN(J,NP).GE.O.DO) V2=V2+DN(J,NP)

IF(DN(J,NP).GE.O.DO) V2J=V2J+DNJ(J,NP)

IF(DABS(DN(J,NP)).GI.DABS(VQM)) VQM=DN(J,NP)

CCNTINUE

PEC(NP)=DCMPLX(V1+V2,V1J+V2J)

REC(NP)=FEC(NP)*CONST

PRINT 19,NP.REC(NP)

PRINT 14,NP.VQM.DN(MAX,NP)

CONTINUE
                                 J=MAX+1-1
600
500
                               CUNTINUE

CONTINUE

VS=DCMPLX(0.D0,1.D0)

DO 800 NP=NS, NUM

REC(NP)=REC(NP)AVS+ANP

IF(IAVE.EQ.0) GG TO 960

SUM=RECO

IF(NS.NE.1) SUM=0.0D0

PRINT 19,NBQ,RECO
 725
800
                                 CALCULATE SUM FOR INTEGRATION CHECK
                               DO 900 I=NS, NUM
PQ=REC(I)
PRINT 19, I, REC(I)
SUM=SUM+2.DO*PQ
PRINT 18, SUM, SUMJ
 900
                              RETURN
CONTINUE
VS=DCMPLX(0.D0,~1.D0)
PRINT 19,NPQ.RECO
SUM=RECO
IF(NS.NE.1) SUM=0.D0
SUMJ=J.D0
DD 950 NP=NS.NUM
VEC(NP)=REC(NP)*DCMPLX(DCOS(NP~RANGL),JSIN(NP*RANGL))
PRINT 19,NP,VEC(NP)
PP=VEC(NP)
SUM=SUM+2.DO*PP
PP=VEC(NP)*VS
SUMJ=SUMJ+PP
CONTINUE
PRINT 18,SUM,SUMJ
KKK=1
                                  RETURN
960
 950
                          RETURN

KKK=1

GO TD 300

FURMAT(1015)

FORMAT(8510.0)

FORMAT(1P8515.6)

FORMAT(1P8515.6)

FORMAT(1P2515.6)

FORMAT(5X, 14,6X, G20.13)

FORMAT(5X, 14,6X, G20.13)

FORMAT(5X, 14,6X, G20.13,6X,G20.13)

FORMAT(5X, 14,6X, G20.13,6X,G20.13)

FORMAT(2X, 13,2X,1P515.6,2X,1P515.6,2X,13,2X,1P515.6,2X

1,1P515.6)

FORMAT(3(1P020.12))

FORMAT(5X,'MAX NOT LARGE ENOUGH ZMAX=',515.8)

285
1234561115
 16
17
18
  10
                                    END
```

END

```
SUBROUTINE GLOS4 (X+A+C+D)
```

GAUSSIAN-LEGENDER INTEGRATION ROUTINE (24 PGINTS)

```
DCUBLE PRECISION DMC.OPC
DGUBLE PRECISION C.D.X( 24).A( 24).XX( 12). AA( 12)
DATA XX(1).AA(1).XX(2).AA(2).XX(5).AA(3).XX(4).AA(4).X
1 XX(6).AA(6).XX(7).AA(7).XX(8).AA(8).XX(9).AA(9).XX(10
2 XX(11).AA(1).XX(12).AA(12)/
* .99516721999702156017999740970-0..12341229799887199540
* .974.25555971309498198391993. -0..28551388628933663181
* .93827455200273275852364900170-0..44277438817419806163
* .8864155270044010342131543419D-0..59298584915430780746
* .8864155270044010342131543419D-0..59298584915430780746
* .8800019859739029219539498726D-0..7334648141080305734
* .74012419157855456424382810510-0..86190161531953275917
* .6480536519369755692524957809D-0..97618652104113888269
* .5454214715885395356583756172D-0..10744427011596563478
* .43379350762604513648708425190-0..115505668055372580135
* .3150426796961653743867932913D-0..12167047292780539120
* .19111886747361536915863982070-0..12583745634682829612
* .6405689286260562608504308262D-1..12793819534675215697
DMC = .5D00*(D+C)
DDC = .5D00*(D+C)
DDC = .5D00*(D+C)
DC = .5D0*(D+C)
DC = .5D0
```

```
SUBROUTINE BESSL (AG. BF. BZ. NUM)
000000000000000
                   SUMPOUTINE RESSE COMPUTES BESSEL FUNCTIONS OF THE RECEIVED ARGUMENT
                 AG = ARGUMENT OF BESSEL FUNCTION
BF = RETURN VECTOR OF DESSEL FUNCTIONS
B7 = 7EKO ORDER BESSEL FUNCTION
NUM = HIGHEST ORDER OF RESSEL FUNCTION TO BE RETURNED
NY*NX = ORDERS CALCULATED BY ASYMPTOTIC EXPANSION
                  IMPRICIT REALER (A-H.O-Z)
DIMENSION BES(1000).6F(1)
                   N7=1000
                   KT=0
                   IF(AG.L7.0) KT=1
IF(AG.LT.0.DO) AG=DA6S(AG)
                  Ny=1F1x(SNGL(AG))
1F(NY.LE.1) NF=16
1F(NY.FQ.2) NF=20
1F(NY.FQ.3)NF=24
1F(NY.GE.4.AND.NY.LT.6) NF=6*NY
1F(NY.GE.6.AND.NY.LT.10) NF=5*NY
1F(NY.GE.10.AND.NY.LT.15) NF=4*NY
1F(NY.GE.15.AND.NY.LT.29) NF=3*NY
1F(NY.GE.29.AND.NY.LT.29) NF=3*NY
1F(NY.GE.80.AND.NY.LT.350) NF=1F1x(1.5*FLOAT(NY))
1F(NY.GE.80.AND.NY.LT.330) NF=1F1x(1.4*FLGAT(NY))
1F(NY.GE.330) NF=1F1X(1.25*FLGAT(NY))
1F(NY.GE.330) NF=1F1X(1.25*FLGAT(NY))
1F(NY.GE.330) NF=1F1X(1.25*FLGAT(NY))
1F(NY.GE.330) NF=1F1X(1.25*FLGAT(NY))
1F(NY.GE.330) NF=1F1X(1.25*FLGAT(NY))
                   NY=YFIX(SNGL(AG))
                   NY=NF
                   NX=11Y-1
                   SET (IRDERS > NY TO ZERO
                   DC 1 I=NY.N7
                   BF3(1)=0.00
CONTINUE
                   COMPUTE ASYMPTOTIC EXPANSION
               DO 2 MU=NX.NY
Z=X/DFLDAT(MU)
AMU=DFLCAT(MU)
A = DSUBT(1.30-24.2)
A1=1.5DU: (DIGG(1.30+A)/Z)-A)
DEL = DEXP(2.DO/3.DO*DLGG(41))
M1 = DEXP(2.DO/3.DO*DLGG(AMU)+DLGG(DEL))
CALL TABLE (X1.AIV.AIPV)
D1 = DEXP(1.DO/3.DO*DLGG(AMU))
D2 = DEXP(5.DO/3.DO*DLGG(AMU))
BG = -5.DO/(43.DO*DLGG(AMU))
BG = -5.DO/(43.DO*DLGG(AMU))
FI = DEXP(0.25DO*DLGG(4.DO*DLCG(AMU))
FI = DEXP(0.25DO*DLGG(4.DO*DLCG(AMU))
BF3(MU) = FI*(AIV/DI+AIPV*BO/D2)
CONTINUE
                   CONTINUE
    2
                   NY2=NY-2
                   USE RECURSIVE EQUATION TO GENERATE ORDERS < NX
                   DO 4 I = 1 + NY2
N=7.Y 2 + I + I
                   BF3(N) = (2.00.0FLOAT(N+1)/X).0F3(N+1)-BF3(N+2)
```

TO THE REPORT OF THE PARTY OF T

SUBROUTING TABLE IS AN AUXILIARY ROUTINE FOR BESSL THAT COMPUTES AIRY FUNCTIONS

AND SERVICES OF SECTION SECTION.

```
IMPLICIT REAL+8 (A-H, N-Z)
DIMENSION AI(50), AIP(50), ABI(15), ABIP(15)
CATA AI/4+0.0D0..562280D0.4*0.0D0..56046200.4*0.0D0.
1.55872400.4*0.0D0..557058D0.4*0.0D0..55545600.4*0.0D0.
2.55351200.4*0.0D0..55242100.4*0.0D0..55098000.4*0.0D0.
3.54458460.4*0.0600..548230D0/
         3.549584E0.4: 0.000.548230D0/
DATA AIP/4.0.0E0.566673D0.4: 0.0D0.569448E0.4*0.0D0.
1.571927E0.4*0.0D0.574520D0.4: 0.0D0.576635D0.4: 0.0D0.
2.576878E0.4: 0.0D0.581056D0.4: 0.0D0.583174D0.4: 0.0D0.
3.585235E0.4: 0.0D0.587245D0/
DATA A61/4: 0.0E0.54823ED0.545636E0.543180D0.
1.540844E0.538618D0.536489D0.534448D0.532488D0.
2.530601D0.528783D0.527027U0/
DATA A81P/4: 0.0D0.507245D0.591120D0.594823D0.
1.598372D0.601782D0.605068D0.608239D0.611305D0.
2.614275E0.617156D0.619954D0/
L=0
            L=0
Y=
             Y = DEXP(DLOG(1.500)-1.500*DLOG(X))
W = DEXP(DLOG(2.00/3.00)+1.500*DLOG(X))
             L = 1
            TF(Y.GT.1.500) GO TO 14
TF(Y.GT.0.49999900) GO TO 1
L=2
             W1 = 10.00 4 Y
             WZ=WI+I.
TWZ=TFTX(SNGL(WZ))
              IZ=IW1+10
              IF(IWI.LT.IW2) IZ=IZ+5
              1Y=17-5
             irīiv.fo.o) ao 10 13
60 10 8
             Y = 10.00 4 Y
1Y=1+1X(SNGL(Y))
1F(1Y-Y) 2.4.6
             12=1Y+1
2
             GO TOB
AIV=AI(IY)
             AIPV=AIP(IY)
             GO TO IS
              17=1Y
              1 Y=1 Y-1
            TY=1Y-1
DP=Y-FLOAT(IY)
IF(L.ED.2) DP=OP/5.00
IF(L.ED.1) GD TO 12
AIV = (1.00-DP):AI(IY)+DP3AI(IZ)
AIPV = (1.00-DP):AIP(IY)+OP3AIP(IZ)
8
             GO TO 19
AIV = (1.DO-GP) ABI(IY)+DP*ABI(IZ)
12
              (11) 91 84:40+(Y1) 918A*(9C-CO-1) = V91A
             gọ to 19
                    = Y/5.DO
13
             DP
             AIV = (1.00-DP) + 0.56419000+DP - AI(IZ)
AIPV = (1.00-DP) + 0.56419000+DP + AIP(IZ)
AIV = DEXP(-0136(2.00)-0.2500*DLGG(X)-W+DLGG(AIV))
AIPV = DFXP(0.5500*DLGG(X)-W+DLGG(AIPV))
AIPV = AIRV - AIRV - 5000
19
              AIPV = AIPV
                                                  (-.5D0)
             RETURN
             PRINT 30
14
30
             FORWAT (5x. 1HAVE EXCEEDED TABLE VALUES!)
             RETURN
                ENO
```

```
SUBROUTINE SUMCK
000000000000
           SUBROUTINE SUMCK COMPUTES THE ANALYTICAL EXPRESSION FOR THE INFINITE SUM OF COEFFICIENTS
           SISUM = INFINITE SUM OF POINT TARGET (IAVE=1) OK
           AVERAGED TARGET (IAVE = 0) CCEFFICIENTS
BKSUM = INFINITE SUM OF BACKGROUND COEFFICIENTS
          IMPLICIT REAL+8 (A-H.G-Z)
CCMMON /OTPRM/ RHG.SIG.WI.H.AL.8E.RX.RY.XO.YO
CCMMON /AVPRN/ NUM.IAVE.PO
COMMON /ANGL/ RANGL
           W=WI
CONST=0.250000
           GAMA=DS ORT (SIG! +2+PO++2)
           1F(1AVE.FO.O) GAMA=SIG
DIV=2.DO.DSURT(2.DO)+GAMA
1F(1AVE.EQ.O) GO TO 10
CCCC
           COMPUTE SUM OF AVERAGED COEFFICIENTS
          A=DERF((2.CO\(X0+RHO)+W)/DIV)
B=DFFF((2.DO\(X0+RHO)+W)/DIV)
C=DERF((2.DDO\(Y0)+H)/DIV)
D=DFRF((2.DDO\(Y0)+H)/DIV)
          $1$UM=CONST:(A-B)$(C-O)
R1=R(W.41.$TG)
K2=R(H.BE.$TG)
CCCC
           COMPUTE SUM OF BACKGROUND COEFFICIENTS
           BKSUM=4.DO'AL+BE#R1+R2/(W+H)++2
           PRINT 5.8KSUM.SISUM
  10
           CONTINUE
           COMPUTE SUM OF POINT TARGET COEFFICIENTS
           A=GERF((2.000*(X0+RHO=DCOS(RANGL)+RX)+W)/DIV)
B=DERF((2.000*(X0+RHO+DCOS(RANGL)+RX)+W)/CIV)
C=DERF((2.00*(Y0+RHO+DSIN(RANGL)+RY)+H)/DIV)
D=DERF((2.00*(Y0+RHO+DSIN(FANGL)+RY)+H)/DIV)
           GO 10 4
           FGRMAT (2x, '8KSJM=', 015, 8, 10x, 'SISUM=', 015, 8)
  5
           END
```

FUNCTION R(H.Z.SIG)

FUNCTION R IS AN AUXILIARY ROUTINE FOR BKSUM

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